Math Camp

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Probability Theory:

- 1) Mathematical model of uncertainty
- 2) Foundation for statistical inference
- 3) Continues our development of key skills
 - Proofs [precision in thinking, useful for formulating arguments]
 - Statistical computing [basis for much of what you'll do in graduate school]

Three parts to our probability model

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1) Sample space: set of all things that could happen

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- 2) Events: subsets of the sample space

Three parts to our probability model

- 1) Sample space: set of all things that could happen
- 2) Events: subsets of the sample space
- 3) Probability: chance of an event

Definition

The sample space as the set of all things that can occur. We will collect all distinct outcomes into the set S

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Examples:

1) House of Representatives: Elections Every 2 Years

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 - $S = \{x : 0 \le x < \infty\}$

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- Key point: this defines all possible realizations

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Notation: x is an "element" of a set E: $x \in E$

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2) Intersection: \cap

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3) Complement of set $E: E^c$

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 - $F^{c} = \{(N, W), (W, W)\}$
 - $S = \Re$ and E = [0, 1]. What is E^c ?

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- $F^c = \{(N, W), (W, W)\}$
- $S = \Re$ and $E = [0, 1]$. What is E^c ?
- What is S^c ? \emptyset

Suppose E = W, F = N. Then $E \cap F = \emptyset$ (there is nothing that lies in both sets)

Definition

Suppose E and F are events. If $E \cap F = \emptyset$ then we'll say E and F are mutually exclusive

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Examples:

- Suppose $S = \{H, T\}$. Then E = H and F = T, then $E \cap F = \emptyset$

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Examples:

- Suppose $S = \{H, T\}$. Then E = H and F = T, then $E \cap F = \emptyset$
- Suppose $S = \{(H, H), (H, T), (T, H), (T, T)\}$. $E = \{(H, H)\}, F = \{(H, H), (T, H)\}$, and $G = \{(H, T), (T, T)\}$

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 - $E \cap G = \emptyset$ $F \cap G = \emptyset$
- Suppose $S = \Re_+$. $E = \{x : x > 10\}$ and $F = \{x : x < 5\}$. Then $E \cap F = \emptyset$.

Definition

Suppose we have events E_1, E_2, \ldots, E_N . Define:

$$\cup_{i=1}^{N} E_i = E_1 \cup E_2 \cup E_3 \cup \ldots \cup E_N$$

 $\cup_{i=1}^{N} E_i$ is the set of outcomes that occur at least once in E_1, \ldots, E_N .

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 $\cup_{i=1}^{N} E_i$ is the set of outcomes that occur at least once in E_1, \ldots, E_N . Define:

$$\bigcap_{i=1}^{N} E_i = E_1 \cap E_2 \cap \ldots \cap E_N$$

 $\bigcap_{i=1}^{N} E_i$ is the set of outcomes that occur in each E_i

Probability

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 - P is a function
 - Domain: all events E

Definition All probability functions, P, satisfy three axioms:

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1) For all events E, $0 \leq P(E) \leq 1$ 2) P(S) = 1

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All probability functions, P, satisfy three axioms:

- 1) For all events E, $0 \le P(E) \le 1$
- 2) P(S) = 1
- 3) For all sequences of mutually exclusive events E_1, E_2, \ldots, E_N (where N can go to infinity)

Definition

All probability functions, P, satisfy three axioms:

- 1) For all events E, $0 \le P(E) \le 1$
- 2) P(S) = 1
- 3) For all sequences of mutually exclusive events E₁, E₂,..., E_N (where N can go to infinity) P (∪^N_{i=1}E_i) = ∑^N_{i=1}P(E_i)

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- Suppose we are flipping a fair coin. Then P(H) = P(T) = 1/2

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- Suppose we are flipping a fair coin. Then P(H) = P(T) = 1/2
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- Suppose we are flipping a fair coin. Then P(H) = P(T) = 1/2
- Suppose we are rolling a six-sided die. Then P(1) = 1/6
- Suppose we are flipping a pair of fair coins. Then P(H, H) = 1/4

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One candidate example:

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One candidate example:

- P(W): probability incumbent wins

One candidate example:

- P(W): probability incumbent wins
- P(N): probability incumbent loses

One candidate example:

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Two candidate example:

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Two candidate example:

- $P(\{W, W\})$: probability both incumbents win

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Two candidate example:

- $P(\{W, W\})$: probability both incumbents win
- $P(\{W, W\}, \{W, N\})$: probability incumbent 1 wins

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Full House example:

- *P*({All Democrats Win}) (Cox, McCubbins (1993, 2005), Party Brand Argument)

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Full House example:

 P({All Democrats Win}) (Cox, McCubbins (1993, 2005), Party Brand Argument)

We'll use data to infer these things

We can derive intuitive properties of probability theory.

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Proposition

 $P(\emptyset) = 0$

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Proof.

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Define $E_1 = S$ and $E_2 = \emptyset$,

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Proposition $P(E) = 1 - P(E^c)$

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Proof.

Note that, $S = E \cup E^c$. And that $E \cap E^c = \emptyset$.

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In words: Probability an outcome in E happens is 1- probability an outcome in E doesn't.

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Proposition

If $E \subset F$ then $P(E) \leq P(F)$.

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We can write $F = E \cup (E^c \cap F)$. (Why?) Further, $(E^c \cap F) \cap E = \emptyset$ Then $P(F) = P(E) + P(E^c \cap F)$ (Done!)

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As you add more "outcomes" to a set, it can't reduce the probability.

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Simulation: use pseudo-random numbers, computers to gain evidence for claim

Tradeoffs:

- Pro Deep understanding of problem, easier than proofs
- Con Never as general, can be deceiving if not done carefully (also, never a monte carlo study that shows a new method is wrong)
- Walk through R code to simulate these two results

To the R code!

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4.2. Three different combination rules were used. We then tried to identify the rules used to combine individual drug predictions into a combination score. Letting P() indicate probability of sensitivity, the rules used are:

$$\begin{array}{rcl} P(TFAC) &=& P(T) + P(F) + P(A) + P(C) - P(T)P(F)P(A)P(C), \\ P(TET) &=& P(ET) = \max[P(E), P(T)], \text{ and} \\ & 5 & 1 \end{array}$$

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Inclusion/Exclusion

Proposition

Suppose E_1, E_2, \ldots, E_n are events. Then

$$P(E_{1} \cup E_{2} \cup \cdots \cup E_{n}) = \sum_{i=1}^{N} P(E_{i}) - \sum_{i_{1} < i_{2}} P(E_{i_{1}} \cap E_{i_{2}}) + \cdots + (-1)^{r+1} \sum_{i_{1} < i_{2} < \cdots < i_{r}} P(E_{i_{1}} \cap E_{i_{2}} \cap \cdots \cap E_{i_{r}}) + \cdots + (-1)^{n+1} P(E_{1} \cap E_{2} \cap \cdots \cap E_{n})$$

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- How many times on the other side? Suppose it appears in m of the E_i m > 0

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count =
$$\binom{m}{1} - \binom{m}{2} + \binom{m}{3} - \dots + (-1)^{m+1} \binom{m}{m}$$

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$$count = \sum_{i=1}^{m} \binom{m}{i} (-1)^{i+1}$$
$$count = -\sum_{i=1}^{m} \binom{m}{i} (-1)^{i}$$

 $\operatorname{count} = -\sum_{i=1}^{m} {m \choose i} (-1)^{i}$

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count
$$= -\sum_{i=1}^{m} {m \choose i} (-1)^i$$

Binomial Theorem: $(x+y)^n = \sum_{i=0}^{n} {n \choose i} (x)^{n-i} y^i$.

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Inclusion/Exclusion

Corollary

Suppose E_1 and E_2 are events. Then

$$P(E_1 \cup E_2) = P(E_1) + P(E_2) - P(E_1 \cap E_2)$$

R Code!

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Proposition Consider events E_1 and E_2 . Then

 $P(E_1 \cap E_2) = P(E_1) - P(E_1 \cap E_2^c)$

Proof.

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$$E_1 = (E_1 \cap E_2) \cup (E_1 \cap E_2^c)$$

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$$P(\bigcup_{i=1}^{N} E_i) \leq \sum_{i=1}^{N} P(E_i)$$

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$$P(\cup_{i=1}^{N} E_i) \leq \sum_{i=1}^{N} P(E_i)$$

Proof.

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$$P(\bigcup_{i=1}^{N} E_i) \leq \sum_{i=1}^{N} P(E_i)$$

Proof.

Proceed by induction. Trivially true for n = 1. Now assume the proposition is true for n = k and consider n = k + 1.

$$P(\bigcup_{i=1}^{N} E_i) \leq \sum_{i=1}^{N} P(E_i)$$

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$$P(\bigcup_{i=1}^{k} E_{i} \cup E_{k+1}) = P(\bigcup_{i=1}^{k} E_{i}) + P(E_{k+1}) - P(\bigcup_{i=1}^{k} E_{i} \cap E_{k+1})$$

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$$P(E_{k+1}) - P(\bigcup_{i=1}^{k} E_{i} \cap E_{k+1}) \le P(E_{k+1})$$

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$$P(\cup_{i=1}^{k} E_i) \leq \sum_{i=1}^{k} P(E_i)$$

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Bonferroni's Inequality

$$P(\bigcap_{i=1}^{n} E_i) \geq 1 - \sum_{i=1}^{n} P(E_i^c)$$

Bonferroni's Inequality

$$P(\cap_{i=1}^{n} E_{i}) \geq 1 - \sum_{i=1}^{n} P(E_{i}^{c})$$

Proof.

 $\cup_{i=1}^{n}E_{i}^{c}=(\cap_{i=1}^{n}E_{i})^{c}.$ So,

Bonferroni's Inequality

$$P(\cap_{i=1}^{n} E_{i}) \geq 1 - \sum_{i=1}^{n} P(E_{i}^{c})$$

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. So,

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Easy Problems

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Probabilistic reasoning pays off for harder problems

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- Assuming leap year counts, N = 367 guarantees at least two people with same birthday (pigeonhole principle)
- For *N* < 367?
- Examine via simulation



Surprising Probability Facts: the E-Harmony Problem

Curse of dimensionality and on-line dating: eHarmony matches you based on compatibility in the most important areas of life - like values, character, intellect, sense of humor, and 25 other dimensions.

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 $\mathsf{Pr}(\mathsf{Exact}) \ = \ \mathsf{Pr}(\mathsf{Agree})_1 \times \mathsf{Pr}(\mathsf{Agree})_2 \times \ldots \times \mathsf{Pr}(\mathsf{Agree})_{29}$

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$$\begin{aligned} \mathsf{Pr}(\mathsf{Exact}) &= & \mathsf{Pr}(\mathsf{Agree})_1 \times \mathsf{Pr}(\mathsf{Agree})_2 \times \ldots \times \mathsf{Pr}(\mathsf{Agree})_{29} \\ &= & 0.5 \times 0.5 \times \ldots \times 0.5 \\ &= & 0.5^{29} \end{aligned}$$

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1 in 536,870,912 people Across many "variables" (events) agreement is harder

Probability Theory

- Today: Introducing probability model
- Conditional probability, Bayes' rule, and independence

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