To Defer or to Differ: Experimental Evidence on the Role of Cash Transfers in Nigerian Couples' Decision-making*

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Abstract

We conduct an original lab-in-the-field experiment on the decision-making process of married couples over the allocation of rival and non-rival household goods. The experiment measures individual preferences over allocations and traces the process of deferral, consultation, communication and accommodation by which couples implement these preferences. We find few differences in individual preferences over allocations of goods. However, wives and husbands have strong preferences over process: women prefer to defer decisions to their husbands even when deferral is costly and is not observed by the husband; men rarely defer under any condition. Our study follows a randomised controlled trial that ended a year earlier and gave large cash transfers over eighteen months to half of the women in the study. We estimate the effect of treatment on the demand for agency among women and find that the receipt of cash transfers does not change women's bargaining process except in a secret condition when the decision to defer is shrouded from her husband. This suggests that the cash transfer to women increases their demand for agency but does not change the intra-household balance of power enough to allow them to express it publicly.

Keywords: intra-household decision-making, bargaining, cash transfers

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1 Introduction

Policymakers focus on progress in women's empowerment and equality because women continue to experience unequal outcomes, even within households (see Duflo, 2012, for a review). Intrahousehold inequality in consumption and health outcomes has been documented worldwide (e.g., Altonji and Blank, 1999; Azmat *et al.*, 2006; Beaman *et al.*, 2017; Bertrand and Hallock, 2001). In addition, large gains in measurable outcomes have been documented for women, children, and even the whole household when women in poor households are provided additional resources such as cash transfers (e.g., Bandiera *et al.*, 2017; Carneiro *et al.*, 2021). Money given to women is more likely to be used for investments in education, children's nutrition, and housing than money in the hands of their husbands (e.g., Adato *et al.*, 2000; Duflo, 2003; Fiszbein *et al.*, 2009; Hoddinott and Haddad, 1995; Thomas, 1994) and increasing women's say in family finances raises savings and investment (Armendáriz and Morduch, 2010; Ashraf, 2009; Duflo, 2012).

Much of the literature assumes that the pattern of unequal consumption is a result of lower empowerment, in line with the dominant economic model of intrahousehold bargaining power in which the utility function of individual members is solely defined over material outcomes (e.g., Browning and Chiappori, 1998; Chiappori, 1988, 1992, 1997; Chiappori *et al.*, 2002; McElroy, 1990; McElroy and Horney, 1981). Indeed, when this assumption is combined with a symmetry assumption about the other-regarding preferences of the spouses,¹ equality in bargaining power implies equality in outcomes. In this model, the welfare weights or bargaining power of spouses is subsumed in a sharing rule which can be inferred from the intrahousehold division of rival consumption (Brown *et al.*, 2021; Cherchye *et al.*, 2017; Dunbar *et al.*, 2013). Implicit in this modelling framework is the assumption that individuals do not care about the way these outcomes are achieved. Perhaps as a result, economists have paid less attention to the *process* of decision–making within the household.

Even when more attention is paid to process, such as whether women have executive agency and whether external interventions can improve female agency within the household (e.g., Dhar *et al.*, 2022; Karimli *et al.*, 2021; Pitt *et al.*, 2006; Riley, 2022), the emphasis remains on the instrumental value of agency, that is, on its capacity to affect material outcomes, rather than on agency's intrinsic value. With a few exceptions (e.g., Afzal *et al.*, 2022; Fernandez *et al.*, 2015),² the simple question of whether empowerment per se is valued by women (similar to the 'capabilities' view of Nussbaum, 2001; Sen, 1999) has been mostly

¹E.g., spouses only derive utility from their own (rival) consumption, or have symmetrical altruistic preferences with identical welfare weights.

²Fernandez *et al.* (2015) examine the correlation between subjective well-being and the right to make particular decisions. Although provocative, the results are not identified.

ignored in the literature on intrahousehold resource allocation.

To examine the link between consumption and empowerment, we present results from an original lab-in-the-field experiment conducted with married couples in Northern Nigeria, a setting within a lower middle-income country where women's agency and rights are severely constrained. We combine traditional elements of economic laboratory design (experimentally assigned controls and separation of individuals) with field elements (a random sample of married women received significant unconditional cash transfers over eighteen months before the study took place) and unique lab-in-the-field elements designed by the team (three laboratory shopping stalls with separate categories of female, male and household items). Finally, the experiment includes a randomly assigned secret treatment in which most decisions are shrouded so neither spouse can tell what their spouse chose or what processes they followed to reach that decision.

We find that the preferences of both male and female participants are egalitarian on average. Despite what seems to be a lack of overt disagreement over budget allocations (as in Almås *et al.*, 2020), spouses have strong preferences over the *way* decisions are made. Women are far more likely than men to consult their spouse, defer to them, and accommodate their wishes. In contrast, male participants tend to make a decision on their own without consulting or accommodating their wife's wishes. Furthermore, we find that women overdefer and men under-defer compared to optimal decisions. The experimental design allows us to conclude that spouses' preferences over the way decisions are made are consequential: they impose an allocative efficiency cost on already very poor households.

The experiment also enables us to estimate the effect of an unconditional cash transfer (UCT) program on both material outcomes and the process by which these outcomes are reached. We can compare the joint decision-making processes of control and treated households, including consultation, communication (honest or not), deferral, revision and accommodation of spouse's preferences and a measure of the willingness to pay for agency over own consumption. From these measures, we can examine several important features of the decision-making process. First, we test, in a controlled environment, whether the randomised cash transfer treatment increases the desire for private consumption and leads to higher material welfare. Second, we test whether treatment has an impact on the willingness to exert agency or to relinquish control to a spouse, either secretly or openly. We do so in a way that enables us to test whether subjects' willingness to pay for agency is driven primarily by instrumental value, or whether subjects deviate from taking the action that would maximise their own material utility. Third, we test whether the treatment increased or decreased efficiency in decision-making.

The UCT intervention had a small effect on the social preferences of men and women;

women allocate more goods or money to themselves if they received the UCT. Treated women are less likely to defer their choice to their husbands, but only when their deferral decision is shrouded and therefore kept secret from their husband. UCT treated women are also slightly less willing to consult. Taken together, these suggest a slight improvement in procedural empowerment but only when decisions are shrouded.

Lab-in-the-field experiments have been used to measure demand and willingness to pay for agency within a household (e.g., Abbink *et al.*, 2020; Afzal *et al.*, 2022; Almås *et al.*, 2018; Iversen *et al.*, 2011; Jakiela and Ozier, 2016; Mani, 2020; Schaner, 2016), but there are only a few studies that look into the effect of a cash transfer on experimental measures of female empowerment. Almås *et al.* (2018), report results from a lab-in-the-field experiment in North Macedonia where a CCT (for children staying in school) was randomly allocated to either the head of household (usually a man) or a woman. They find that, compared to wives of male recipients, female recipients of this long-term CCT exhibit a lower willingness to pay (WTP) to appropriate an additional windfall for themselves instead of their spouses. This, they argue, shows an improved bargaining power of the CCT-receiving women: they are willing to pay less than others because they have more say over the allocation of the windfall even if it is received by their husband. However, other channels could explain this finding³ and our study, by including a secret condition treatment, helps unwrap some of the issues raised in that paper.

In the following section, we develop a conceptual framework for analyzing the behaviour of our participants. Section 3 explores the setting and details of our lab experiment and explain how we use it for examining the behaviour of households. Our empirical work is divided into two sections. First, in Section 4 we examine the behaviour of the households in the control group, a baseline of behaviour. Second, in Section 5 we examine how this behaviour was changed by the UCT experiment. Finally, in Section 6, we conclude.

2 Conceptual framework

In the canonical model of intrahousehold consumption allocation (e.g., Becker, 1964, 1981; Chiappori, 1988; McElroy and Horney, 1981), the decision that a couple makes is represented

³For example, fairness norms could explain why women who were beneficiaries of a CCT program are less likely to object to a new transfer to be given to their spouse instead of receiving it themselves again. Moreover, former recipients of a CCT might want to avoid 'taxation' from family and relatives (see Jakiela and Ozier, 2016) after experiencing it first-hand and may rather have their husbands deal with such attempts. This issue is salient in Almås *et al.* (2018) because the information on being a recipient of the CCT program was common knowledge.

as the solution to:

$$\operatorname{Max}_{x} \omega u_{h}(c_{h}) + (1 - \omega)u_{w}(c_{w}) \qquad \text{subject to } c_{h} + c_{w} = y \tag{1}$$

where w and h represent the wife and husband, respectively, y is the combined income of the couple and c_h and c_w are the consumption expenditures of the husband and wife, respectively. The relative welfare weights of the husband and wife, ω and $1-\omega$, represent bargaining power over consumption: the larger ω is, the larger c_h is relative to c_w . It follows that the relative bargaining power of the two spouses can be inferred from their consumption: an empowered wife consumes more (Cherchye *et al.*, 2017).⁴

A simple theory of intrahousehold agency Empowerment, however, is also about having a say in decisions that affect you. The call for female empowerment is partly because of its anticipated material benefits, as it allows women to better defend and promote their interests. But it is also about the process by which decisions are made: just as a country is not considered democratic simply because it fulfills the material needs of its population, the empowerment of women cannot be solely judged by their material welfare. Which spouse believes they have the right to decide how to allocate the household's consumption budget? Who consults and who is consulted on that decision? Does the consulted spouse distort his or her reported preferences? Are reported preferences taken into account by the other spouse? Our experiment is designed to capture each of these important dimensions of the decision–making process within the couple while controlling for the consumption preferences of each spouse.

Afzal *et al.* (2022) offer a straightforward formalization of preferences over process (see also Doepke and Tertilt, 2019). Let there be two spouses h and w and let c_h and c_w be their consumption bundles. According to their definition, spouse w has a preference for executive agency on her consumption if:

$$u_w(c_w, D_w^h) < u_w(c_w, D_w^w) \tag{2}$$

where consumption c_w is kept constant, D_w^h means that spouse h decides the consumption bundle c_w for w, and D_w^w means w decides for herself. The above inequality implies that when w (or h) prefers to decide, she is willing to accept a lower level of material welfare in exchange for having executive agency on her consumption.

⁴This logic is behind a series of recent papers on Engel curves within households (e.g., Bargain and Donni, 2012; Bargain *et al.*, 2014; Brown *et al.*, 2021; Calvi, 2020; Dunbar *et al.*, 2013; Lechene *et al.*, 2022; Sokullu and Valente, 2022; Tommasi, 2019).

In this paper we do two things: 1) we experimentally assess the *decision process* by which couples allocate consumption between themselves and estimate spouses' demand for control and agency within that process; and 2) we estimate the causal effect that a full-scale external intervention — an unconditional cash transfer (UCT) — has on the decision process and on the demand for agency of each spouse. We know, from an evaluation of the UCT, that this UCT increased women's self-employment and income and that it raised traditional indicators of female empowerment focusing on material welfare and on decision–making in farm production (Papineni *et al.*, 2024). We want to know whether it also affected agency and, if yes, along which dimensions.

Our experiment starts by eliciting true allocative preferences in consumption, which we identify by letting each spouse independently decide how to allocate various consumption budgets y between c_w and c_h , as well as between individual and joint consumption. This is achieved by setting up stalls that carry gender-specific goods. One advantage of this approach is that it does not assume that each spouse is solely motivated by self-interest: it allows subjects to manifest (possibly different) altruistic preferences in household consumption.⁵ This differs significantly from the canonical model of intrahousehold allocation in which differences in material allocation between spouses are seen as a manifestation of different bargaining weights. When spouses are differentially altruistic towards each other, the canonical model cannot identify how power is distributed within the household. Furthermore, with altruistic preferences, it is possible for both spouses to divide a budget y in the same manner $c_w + c_h = y$ while at the same time both wanting to make that decision themselves:

$$u_w(c_w, c_h, D_w^h) < u_w(c_w, c_h, D_w^w) \text{ and } u_h(c_w, c_h, D_h^w) < u_h(c_w, c_h, D_h^h)$$
(3)

We experimentally measure this demand for executive agency by allowing subjects to *defer* to their spouse the division of a budget y between consumption bundles c_w and c_h . Deferral refers to a subject's willingness to delegate consumption decisions. It indicates a reluctance to hold sole *executive agency* over household consumption decisions. In our model, this preference can be represented by reversing the sign of equation (2):

$$u_w(c_w, c_h, D_w^h) > u_w(c_w, c_h, D_w^w)$$

$$\tag{4}$$

Demand for executive agency can be measured experimentally: we let subjects independently and secretly decide how to divide a budget y into c_w and c_h on their own and then we ask

⁵Altruistic behaviour can arise from a variety of sources, such as social norms, moral and philosophical principles, affection, or reaction to emotional stimulus. Our experiment is not designed to disentangle them.

them whether they would prefer to delegate that division to their spouse. We can also look at whether *both* husband and wife delegate the decision to their spouse, suggesting a joint preference towards a negotiated outcome.

Delegating the final decision to someone else does not, however, imply an absence of demand for agency: the person may still wish to *influence* that decision, e.g., because of asymmetric information about preferences between spouse (e.g., Tagat *et al.*, 2023) or because agency is valuable in its own right (e.g., Afzal *et al.*, 2022).

Measuring demand for consultative agency is more complex because consultation involves at minimum three distinct steps: truthful communication of a preferred split $(c_w, c_h)^w$ by w to h; consultation by h of this communication made by w; and accommodation of that communication by h who revises his original division budget allocation $(c_w, c_h)^h$ at least partially in the direction of the split communicated by w.

To capture these ideas, we invite subjects to communicate a preferred allocation to their spouse. Ideally, this allocation should correspond to their true preferred allocation, in case the spouse decides to implement it. It also may involve some *strategic misrepresentation* whereby the subject distorts their true preferences when communicating with their spouse, e.g., to avoid disapproval or retaliation. Misrepresentation may also indicate an unwillingness to reveal one's true preferences due to self-image considerations (e.g., not wanting to appear too selfish). The absence of misrepresentation therefore signals a relatively serene approach to joint decision–making among spouses and is a pre-condition for consultative agency to achieve allocative efficiency in consumption.

Communicating one's preferences to a spouse may fall on a deaf ear, though: the spouse may refuse to hear them. Willingness to listen to someone else's preferences is what we call consultation. Two-way communication between the spouses is thus a necessary condition for a subject to influence their spouse's decision. But it is not sufficient: the communicated preferences must also be taken into account by the deciding spouse, i.e., they must be accommodated. Taken together, truthful communication, consultation, and accommodation therefore create the conditions for a negotiated allocation of the household consumption budget. Even in the presence of deferral, the combination of these three processes confers to the deferring spouse the right to be heard and influence decisions — a process we call *consultative agency*. It is measured as the combined effect of communication (including misrepresentation) by the subject and of their spouse's willingness to consult and accommodate.

Formally, let c_w^w be the consumption for self, chosen by subject w in isolation; let c_w^h be the consumption for w chosen by subject h in isolation; let $k_h = 1$ be h's decision to consult w's communicated preferences and 0 otherwise; and let $c_w^{h_accom}$ be the consumption for wchosen by h after being told w's communicated preference. Consultative agency is then defined as the extent to which c_w^h moves towards c_w^w as a result of the consultation process. It is given by the following formula:

$$C_w^h = 1 - \frac{c_w^w - (k_h c_w^{h_accom} + (1 - k_h)c_w^h)}{c_w^w - c_w^h}$$
(5)

where we have normalized by $c_w^w - c_w^h$, the difference between the true preferences of w and h. Note, first, that C_w^h is undefined if $c_w^w = c_w^h$; if the husband and wife both prefer the same amount to for the wife, we can't observe accommodation. Whenever the initial preferences are different, however, $C_w^h = 1$ means that w receives her true preference while 0 means that she receives her husband's preference. Any value between 0 and 1 indicates a compromise between the spouses. It is also possible for $C_w^h < 0$, in case h revises his initial choice c_w^h away from c_w^w , for instance, to manifest disapproval, or for $C_w^h > 1$ in case h overcompensates by moving beyond what w asked for. To facilitate the analysis, we constrain C_w^h to only take values between 0 and 1.

This expression can be decomposed into those parts which are due to misrepresentation, willingness to consult, and extent of accommodation:

$$C_w^h = 1 - \left(\frac{c_w^w - c_w^{w.comm}}{c_w^w - c_w^h} + \frac{(1 - k_h)(c_w^{w.comm} - c_w^h)}{c_w^w - c_w^h} + \frac{k_h(c_w^{w.comm} - c_w^{h.accom})}{c_w^w - c_w^h}\right) \tag{6}$$

where c_w^{comm} is the choice communicated by w to her husband. The first fractional term is the loss of agency due to misrepresentation by w, the second fractional term is the loss of agency due to lack of consultation by h, and the last fractional term is the loss of agency due to the lack of accommodation. An equivalent formula can be derived for the consultative agency of the husband – should the wife be granted sole executive agency.

We also borrow aspects of the design developed by Afzal *et al.* (2022) to study the decision to defer when it has no instrumental value. This allows us to parse out the role played by instrumental motives from preferences over process — a distinction that mirrors that between consequentialist/utilitarian and deontological ethics (e.g., Benabou *et al.*, 2024). In this part of the experiment, w chooses one of four physical consumption bundles for herself to consume on the spot. She is then invited to pick a bundle for h without knowing which bundle h prefers and is finally asked whether she wants to defer the choice to h. Mirror decisions are independently made by h. Here, *not* deferring the choice to the spouse has no instrumental value — w cannot appropriate any of h's bundle and vice versa. It is therefore always inefficient since choosing for the spouse has a strong chance of deviating from his or her preferred bundle. Hence choosing *not* to defer is a measure of demand for pure control (e.g. Afzal *et al.*, 2022; Ashraf, 2009; Jakiela and Ozier, $2016)^6$.

Finally, we examine how spouses handle the trade-off between efficiency and self-interest in household production decisions, using a simplified version of a game developed by Fafchamps and Kebede (2022). The literature has shown that assigning (legal or customary) control over financial or productive assets within the couple can affect the allocation of consumption expenditures between spouses (e.g., Browning *et al.*, 1994; Lundberg *et al.*, 1997). In some cases, this can even lead to inefficient choices (e.g., Lundberg and Pollak, 1993; Udry, 1996). The purpose of this part of the experimental is to test whether subjects are capable of taking efficient production decisions even if it means reallocating productive endowments to their spouse. Deviation from the efficient choice again captures demand for pure control, in this case over the production process.

A simple theory of emancipation Our experiment, by combining these different elements, produces evidence on how executive agency is allocated between spouses and the extent to which spouses are granted consultative agency in the absence of external intervention. We then use the experiment to test how this equilibrium is affected by two separate interventions: secrecy in executive agency, which is granted to some subjects in the experiment; and a large UCT intervention targeted to the wife. The objective of these interventions is to disentangle three broad mechanisms of subjugation which we refer to as oppression, indoctrination, and alienation.

Oppression refers to the fear of retaliation: a spouse may wish to exert control over the allocation of household consumption, but fears disapproval or reprisal if this action is observable by the spouse (e.g., Buchmann, 2022). Secrecy may also enable individuals to insulate their income from social pressures from the spouse (e.g., Zhang, 2023a,b) or from outside the household (e.g., Boltz *et al.*, 2019). By providing the opportunity for a subject to exert executive agency in secret (e.g., Ashraf, 2009), we elicit what we call 'pent-up demand' for agency, that is, demand for executive agency that is repressed by the spouse and can only be expressed in secret (e.g., Ashraf *et al.*, 2014). Reprisal may also take a financial form, e.g., by diverting household resources away from the offending spouse. In this case, providing more financial autonomy to an oppressed spouse should offer protection against such reprisal. Hence an intervention that helps that spouse gain financial independence such as the UCT (see Papineni *et al.*, 2024) — should reduce the fear of reprisal, thereby allowing the spouse to openly exert more agency, either executive or consultative, without the need for secrecy.

⁶For a deeper understanding of the difference between instrumental and intrinsic control, see the following papers on intimate partner violence (IPV) (e.g., Angelucci, 2008; Bobonis *et al.*, 2013; Haushofer and Shapiro, 2016; Hidrobo *et al.*, 2016; Perova and Vakis, 2013; Roy *et al.*, 2018).

Indoctrination refers to the internalisation of gender norms. Other researchers have sought to document the effect of a change in such norms — or in the perception of such norms — on behaviour and agency (Bursztyn *et al.*, 2020). In this paper, we do not attempt to change norms or perceptions directly. Nonetheless we are hopeful that gender norms may improve in the aftermath of a sustained intervention that, by providing regular monetary transfers to married women, helped them start a business and gain financial autonomy (Papineni *et al.*, 2024). If this is the case, we expect the UCT intervention to modify the household budget allocation in favour of women, to induce husbands to grant more consultative agency to their wife — i.e., by consulting them and accommodating their stated preferences — and to induce wives to be less willing to defer consumption decisions to their husband and more willing to truthfully disclose their true preferences to their husband.

The third main channel for subjugation is alienation, which often manifests itself through a lack of aspiration for agency: women do not imagine that they could make decisions on their own (e.g., Bernard *et al.*, 2023; Orkin *et al.*, 2023). In individualistic societies, executive agency over own consumption is taken for granted. But in collectively-minded societies (e.g., Enke, 2019; Roland, 2020), making important decisions without consulting with the group is often regarded as unwise or illegitimate. Consulting others and deferring to those in charge of the group is regarded as beneficial to the group and its members — and in patriarchal societies the application of this principle to women and other dependents means that they do not imagine being capable of making correct decisions by themselves. This makes it all the more important to capture consultative agency in contexts, such as our study, where collective decision–making is generally seen as the norm (Thomas *et al.*, 2023).

If women refrain from expressing a demand for agency because of collective thinking, the decisions they take should be the same whether they are taken in secret or not: there should be no pent-up demand for agency that is expressed in secret. An external intervention that specifically targets funds to one member of the household does, however, create a new reality in which the individual agency of married women is recognised by an external actor. This in turn may favor a rise in individualism that would express itself in more selfish consumption choices and a demand for executive agency. If such demand may trigger reprisal from the spouse or the social group, it may only be expressed in secret.

3 Experimental Design and Background

We conducted a lab-in-the-field experiment in 38 sessions with married couples in 27 villages from two rural regions in Kebbi State, in northwest Nigeria. The experiment was embedded in a larger randomised controlled trial of an unconditional cash transfer (UCT) program targeted at ultra-poor women in this region. The majority of households in this region live in extreme poverty. These households are predominantly of Hausa ethnicity and Islamic faith, known for their patriarchal social structures and women face restrictive social norms, low levels of education, and low labor force participation, resulting in limited agency (Braimah, 2014; Ogu *et al.*, 2016).

The UCT baseline survey confirms women's limited agency, especially regarding labor outside the home which is decided on by husbands, and it documents the existence of social norms restricting women's work and mobility, with 40% of men considering it inappropriate for women to accept paid jobs outside the home. Female labour force participation was low before the intervention, with most women engaged in household work or childcare. Only 36% of women reported any income-generating activities in the twelve months prior the baseline survey, mostly on household-operated farms (Papineni *et al.*, 2024).

The UCT, randomised at the household level within study villages, provided a total of 75,000 Nigerian Naira (roughly USD 693 PPP for 2015) over fifteen months to the primary female decision-maker of the household. This is a considerable amount of money for the participating households, amounting to approximately half of their annual consumption. Indeed 92% of sample households lived on less than USD 1.90 a day according to the baseline survey.

The cash transfers did not come with any explicit conditions of how the money should be spent or shared. However, during a sensitisation campaign, households were told by traditional community leaders that the money was for the female recipient. As noted by Papineni *et al.* (2024), this light intervention could have influenced the female recipient's ability to keep more of the cash transfer (e.g., Benhassine *et al.*, 2015). Bastian *et al.* (2017) report that 54% of the cash transfer (in the same program) was kept by female recipient and 26% was passed on to her husband; per capita consumption increased by 25%; and the value of household assets (mainly small animals) increased by 30%. Papineni *et al.* (2024), similarly, show an increase in a women's empowerment index as measured by a modified versions of the Women's Empowerment in Agriculture Index (WEAI).

The experimental protocol is detailed in the rest of this section. Further details are available in Section C of the Online Appendix.

3.1 Budget allocation decisions

At the heart of our design is a set of budget allocation decisions that subjects are asked to make under several treatment conditions. In the first part of the experiment, each spouse independently and secretly chooses how to split a budget of 2500 Naira (around 20 USD) between two categories of goods.⁷ This choice is never revealed to their spouse but it has a positive probability of being implemented, making revealed preferences incentive-compatible. These choices can be seen as defining the social utility of each spouse if he/she were given full control over household consumption — i.e., they measure $W_i(x)$ for a particular x vector and fixed budget. y = 2500.

For three of the four domains (A, B and C) subjects are given a budget and asked how they wish to allocate it to pre-stipulated categories of goods they can only purchase from us at the end of the experiment, thereby reducing the chance of ex-post transfers outside the experiment. Once that budget has been assigned to a particular category, it cannot be converted into cash or spent in another category. To facilitate comparison with other experiments in which spouses are given cash, we also include a separate allocation decision (Domain D) in which subjects divide a cash amount between themselves and their spouse.

To implement this design, three market stalls are constructed in the lab, each of which contains only one category of items. These items were pre-tested to be recognisable and desirable, and they include some goods that are easy to buy locally and some goods that only available in town. All participants are shown pictures of the items available in each stall. One stall includes goods intended to appeal to women: jewellery, colorful fabric, and dresses. The second stall includes male-oriented goods such as hats, caps, belts, shoes, and fabric. The third stall offers common household items such as cleaning supplies, mats, cups, plates, and mosquito coils.⁸ The choice of these items was carefully designed such that male goods would not appeal to women and vice versa. Given how strongly gendered the local context is, this was not particularly difficult to do.

At the end of the experiment, husband and wife are reunited and, if that decision was randomly selected for implementation, couples are given vouchers redeemable in specific stalls. Thus, when a subject allocates funds to a stall expected to appeal to their spouse, the subject may intend to purchase something for their spouse or to let their spouse choose – in the context of this experiment, both choices are observationally equivalent. We did not record the items that subjects purchased or who made the choice.

By making two of the stalls gender-specific, the design limits the scope for ex-post reallocation of the objects between spouses. There still remains the possibility that subjects sell or give the objects to others after the experiment. To avoid this, we include a fifth domain in which subjects separately consume a food and drink of their choice directly in the lab, as in Afzal *et al.* (2022). In that case, reallocation across spouses is impossible.

⁷This amount is roughly equivalent to six days of household consumption or around two weeks of the UCT payments. It was chosen to be sufficiently large to incentivise participants to take the experimental tasks seriously.

⁸See Online Appendix photographs C3, C4 and C5, respectively.

Throughout the experiment, each spouse answers various allocation decisions regarding their preferred budgetary allocations over four possible choice pairs:

- A female vs male items purchased in the lab stalls
- B household items vs male items purchased in the lab stalls
- C household items vs female items purchased in the lab stalls
- D cash for wife vs cash for husband⁹
- J two kinds of cookies and two flavours of juice for both self and spouse.

Choice A measures the social utility $W_i(x_w, x_h)$ of consumption bundles x_w and x_h for individuals $i = \{w, h\}$, conditional on the choice of goods available in our stalls.

Choice D is over money and can thus potentially be undone outside the lab, but it offers the advantage that consumption is not restricted to the goods on sale in the lab. It measures the utility $W_i(x_w, x_h)$ of consumption budgets x_w and x_h for individuals $i = \{w, h\}$ that can be achieved from the choice of goods available outside the lab.

Choices B and C are between private goods and household goods, thereby revealing how husbands and wives differ in the extent to which they care for household public goods, including child consumption (x_c) . They measure $W_i(x_w, x_c)$ and $W_i(x_h, x_c)$, respectively.

In each case A-D, the subject is given a budget to divide — in multiples of 100 Naira — between each of the two options listed above. In the rest of the paper, we refer to each of these choice sets using the letters A to D above. The food-and-drink game is only played once with real stakes — i.e., actual food and drinks given to the subject at the end of the experiment. At the end of the experiment, subjects are also asked to choose between two input allocations that determine individual incomes — see below.

3.2 Decision blocks

The experiment is divided into a series of decisions grouped into blocks that each subject makes silently and in isolation from their spouse — who is in another room. In terms of sequencing, Block 1 always comes first while Blocks 4, 5 and 6 always come last, in that order. The order of Blocks 2 and 3 is permuted at random across sessions.

At the end of the experiment, one decision from one of Block 1 through 4 and 6 is selected at random for each couple to determine the cash and vouchers received by the couple. This

⁹The decision was presented as money for self or money for spouse but we flip the husbands decisions and always present it as money for wife versus money for husband.

means that each decision a subject makes is fully incentivised. 1 provides examples of the flow of decisions across blocks 1 to 4 and how final payoffs are determined. We delineate which choices were secret for everyone and which ones were secret only in the secret treatment, and, in the next sub-section, explain how choices were hidden.

Block 1 Allocating a budget across different types of items or payments [split and resplit]: Each participant is sequentially asked to split a budget of 2500 Naira (about \$23) in 100 Naira increments. This is done in each of the four domains: splitA-D. Each subject is then asked to split a budget of 2100 Naira¹⁰ in domains A and D: resplitA and resplitD. The order of both sets of choices is the same for all subjects in a session and randomised across sessions. These choices are never revealed to the spouse. The purpose of this block is to elicit each subject's true preferences over all possible choice pairs.

Block 2 Choosing whether to allow one's spouse to make the decision instead of making one's own [defer]: Subjects are reminded of their initial allocation across splitA-D and can choose to retain this allocation or to replace it with their spouse's allocation, which they have not seen (deferA-D). The object is to elicit subjects' willingness to defer budget allocation to their spouse. This decision is not revealed to the spouse in the secret treatment.

Block 3 Choosing whether to allow one's spouse to make the decision instead of making one's own with additional costs [defercost, deferbenefit]: Having already made decisions in Block 1, subjects are first reminded of their initial allocation across splitA and splitD and then given a choice between their decision over a budget of 2500 or their spouse's decision over a budget of 2100 (defercostA and defercostD). In addition, subjects are given a choice between their decision applies only to domain D, splitting money between husband and wife (deferbenefitD). For defercost it is costly to defer the decision and for deferbenefit it is costly to retain the decision. The purpose of this block is to elicit subjects' willingness to pay to defer — or not defer — budget allocation to their spouse. These choices are not shown to the spouse in the secret treatment.

Block 4 Communicating preferences, consulting over preferences and revising decisions [communicateB-D, consultB-D, reviseB-D] Subjects are then asked to pick an allocation across domains B, C and D that will be communicated to their spouse communicateB-D). The purpose of this part of the experiment is to determine whether subjects choose to misrepresent their true choices from their spouse. These are the only decisions in the secret treatment that are always visible to the spouse without shrouding.

 $^{^{10}}$ The difference of 400 Naira was chosen so as to be just large enough to identify variation in deferral rates among women. In the last two sessions (10% of subjects), we increased the difference by reducing the amount to 1800 Naira to see whether we could induce deferral among men.

After having done this, subjects are asked if they would like to see the communication of their spouse (**consult**B-D). The purpose of this question is to ascertain whether subjects are interested in learning about their spouse's preferences — something they would want to do in case they are willing to accommodate these preferences, in part or in full.

Finally, whether or not the subject chose to see their spouse's communication, the subject is shown the communication¹¹ of their spouse and asked if they would like to revise their split decisions made in Block 1. It is indeed possible that a subject does not wish to be informed of his or her spouse's preference but, once informed, cannot resist the mental pressure of accommodating their wishes. This phenomenon, if observed, would be reminiscent of the findings of DellaVigna *et al.* (2012) whose subjects avoid being asked for something they have difficulty refusing when asked. Variable **revise**B-D records the revised decision after receiving communication: if a subject did not want to change their split, **revise** is the same as the original **split**. The purpose of this part of the experiment is to determine the extent to which each subject accommodates the revealed preferences of their spouse.

Block 5 Food and drink choices [deferF] In this block, each subject is asked to select one of two different cookies (Food) and one of two different drinks (Juice) for themselves and for their spouse. Having chosen for themselves and their spouse each is given the option to defer their decision over both food and juice to their spouse (deferF) without knowing what their spouse chose for them. The purpose of this question is to elicit the subject's willingness to defer their private consumption of food and drink to their spouse even when that own consumption has no effect on the spouse's own consumption and their consumption is not observed by the spouse, as in Afzal *et al.* (2022). Subjects who do not defer receive the food and drink of their choice; those who defer receive the food and drink chosen for them by their spouse. The food and drink are consumed privately by each subject in their genderspecific room so the spouse does not affect own consumption.

In about a third of the sessions (13 sessions with 166 subjects), there is a penalty for retaining the decision: the subject receives one cookie and a half glass of juice of their own selection if they do not defer; but they receive two of cookies and a full glass of the options that their spouse chose for them if they do defer. The purpose of this design is to elicit a bound on subjects' willingness to pay to defer their own consumption choice to their spouse.

Table 1 shows two examples of how decisions are made and paid out in Blocks 1 through 5 for domain B (male vs. household goods). Note that the first woman chooses to spend 1400 on male goods, communicates to her spouse that she wants to spend 1200, asks to see her spouse's preferences, and then moves towards his communication. The second woman

¹¹All subjects retained the right not to look at the information they were given.

chooses to spend 1300 on male goods, communicates her true preferences, does not want to see her husband's preferences and, when shown them anyway, does not alter her decision. Note that, in both examples, by chance, the husband's decision is implemented, but the first couple had the woman's choice to defer selected, and the second couple had the man's original decision selected. In Block 5, the first woman wanted round cookies and coke, guessed what her husband wanted and then chose to consume what he wanted for her (not shown). In the second example, she chose to consume what he chose for her.

Block 6 Allocating inputs across two production functions [efficiency] In this final block, the focus of the choice is not consumption but production. Each subject is asked to pick one of four possible input allocations between themselves and their spouse. Each input allocation maps into a cash income for themselves and a cash income for the spouse.¹² Subjects answer two versions of the game.¹³ Each game has an efficient choice that maximises the total income of the household. In Game 1, the efficient choice implies less output is assigned to self and in Game 2, the efficient outcome assigns less to the spouse.

3.3 The secret treatment

We randomised the sessions equally into a secret treatment in which most decisions are shrouded and a no-secret treatment in which most decisions are not shrouded. Other than Block 5 (cookie and juice), which is always implemented during the lab experiment, all other choices are only implemented if selected at the end. Some decisions are always shrouded from spouses, irrespective of the secret condition. In particular, all Block 1 decisions (initial allocations) are always hidden from the spouse. In contrast, the communication sent to the spouse in Block 4 and 5 is never hidden. All the other decisions, if selected at the end of the experiment, are either potentially revealed or shrouded depending on the secret condition.

Shrouding of these decisions is achieved as follows. At the end of the experiment, when a decision is selected for implementation, if the decision is to be shrouded, the actual decision is first mixed with a random decision, one of the two is drawn at random and then implemented. This ensures credible deniability in the secret condition: the spouse never knows whether the implemented allocation was randomly selected or selected by the subject. The precise process is described in further detail in Section C4 of the Online Appendix.

Once this process is over, for domains A, B and C, the selected individual receives tokens to be spent in the stalls, and the couple is brought together to the lab stalls associated with that decision. For domain D and Block 6 (the production game), the money is split between

 $^{^{12}}$ The games mimic Udry (1996) in which men and women farm separate plots and fertilizer should be allocated across the plots to achieve the maximum level of output.

¹³For a small number of sessions, only one version was presented to subjects.

the husband and wife according to the choices (or shrouded choices) of the individual whose decision was selected to be implemented. Block 5 is always selected and is implemented (either the choice or the decision to defer) privately in the gender-separated room during the experiment.

3.4 Implementation

The UCT randomised controlled trial (RCT) was designed by the Africa Gender Innovation Lab of the World Bank. The UCT intervention, supported by USAID and implemented by the Catholic Relief Services, took place between September 2015 and March 2017. The baseline survey for the RCT took place between April and June 2015. A midline survey was conducted approximately one month after the last UCT monthly payment, i.e., between April and June 2017. The endline survey was conducted one year after the end of treatment, in May to July 2018.

The lab-in-the-field experimental sessions took place in 27 villages from the UCT impact evaluation sample in 38 distinct lab sessions between March and May 2018, one year after the UCT intervention ended and just before the endline survey began.¹⁴ The one-year gap between the RCT and our experiment is intended to diminish the pure income effects of cash receipts to identify longer-term effects on decision–making within the household. To ensure that participants could easily walk to a location that allowed sufficient privacy, we selected villages from the evaluation sample with a nearby school. This means that our sample is not necessarily representative of the full UCT study sample. Lab sessions were conducted either on off days or after regular school hours.

Participants for the lab sessions were recruited among married women included in the UCT experiment either as control or as treated. Female participants were invited to participate to the lab experiment with their husband. Upon arrival, participants were told that they would receive a compensation of 500 Naira (around \$5) as show-up fee. Additionally, they were informed that they could earn significantly more, either in cash or in-kind, based on their decisions in the experimental games, in the sense that one of all the decisions made by both husband and wife that day would be selected for implementation and additional pay-out. Husbands and wives were then sent to two separate gender-specific rooms where the bulk of the experiment took place. The female room only included female enumerators; the male room only had male enumerators. As is common in settings where illiteracy is widespread, all instructions were read to subjects by a room supervisor or an enumerator

¹⁴Nine of them took place in villages that received an intervention from the Feed the Future Nigeria Livelihoods Project (FNLP). This intervention is modeled on an ultra-poor graduation program focusing on focuses on education, coaching and savings groups – but without monetary transfers.

who sat with each subject and asked them to make decisions by pointing to pictures using experimental prompts. This setup allowed us to reach a wide range of subjects and avoided participants revealing their choices to others verbally. When allocating money to different decisions or budgets, participants used laminated photocopies of Nigerian Naira denominations with which they were familiar. Participants spent approximately 2.5 to 3 hours in the lab session and earned up to an equivalent of 3000 Naira (around 25 USD) – 500 as show-up fee and up to 2500 in cash or in kind, depending on the random draw.

Recruitment was stratified to ensure that half of the lab participants had received the UCT and the other half had not. Appendix Table A1 shows that the lab sample is balanced on observables across UCT treated and controls.¹⁵ It also indicates that, on average, women in the lab sample are about ten years younger than their husbands and are unlikely to be considered adequately empowered at baseline, according to various empowerment indicators, including the A-WEAI.

In terms of power, we have 506 participants in 38 lab sessions and four equally-sized treatment cells resulting from the crossing of the UCT and the secret condition. Given that we cluster all standard errors at the level of the lab sessions, we are powered enough to identify large effects of combinations of the UCT intervention and the secret condition: the minimum effect size to reject the null with 80% probability is 14 percentage points in all the dichotomous choices – such as the decision to defer or consult. Given that the UCT was found to have effects exceeding 25 pp on many outcomes, including empowerment (Papineni *et al.*, 2024), we can reasonably expect causal effects of a similar magnitude in our experiment. For continuous variables such as split choices, these decisions are typically unaffected by the secret condition. For these, the minimum effect size that we can reject with large probability is 85 Naira – which is small (3.4%) relative to the choice budget of 2500 Naira.

4 Preferences over outcomes and process in the UCT control group

To properly appreciate the impact of the UCT intervention, we examine, first, the behavioural patterns of experimental participants in the control group, that is, those who did not receive the UCT treatment.

¹⁵The test of balance between secret and no secret treatments also shows no significant differences. A test across all four treatment cells (not reported) shows that men and women in the UCT+secret treatment cell are 4 and 2 years older than average, respectively. All other variables are balanced.

4.1 Budget allocation decisions

We present, in Table 3, summary statistics of all the key behavioural variables, broken down by gender. A t-statistic for the test that the means (pairwise within couples) are equal is also provided, together with the associated p-value.

The first panel of Table 3 shows the four main split decisions with the full budget of 2500 Naira. Most split decisions are significantly different by gender, but the average differences between husbands and wives are not large in magnitude. Both spouses tend to divide budgets more or less equally on average, and this is true for all four splitting decisions (1250 would be an equal split) — a finding reminiscent of equal sharing in dictator games. However, these averages hide a lot of variation across the sample with significant proportions of both men and women allocating much less or much more to themselves.¹⁶ Note that men tend to allocate a larger budget share to female goods sold in the lab (1298>1250), but less to money for their wives (1062<1250); money that could spent outside the lab. Note that if we convert the amount a man gives to his wife in goods or money into the amount he keeps for himself, women are more selfish in goods than men are (1410 > 1202, p-value<0.00) and women are less selfish in money than men are (1265 < 1438, p-value<0.00). This may reflect the fact that the items sold in the lab shop appeal more to women than men or that men had planned to buy female goods on their wife's behalf both in and outside of the lab.¹⁷

Given that these splitting decisions were always shrouded, the findings violate the idea that, given the opportunity, spouses would like to appropriate a large share of the offered budget to themselves. This is true of women but also of men who, as we shall see shortly, wield most of the power in our sample population. From this evidence we conclude that spouses have social preferences, meaning that they incorporate the expenditures of each other in their own utility function. We also note that women do not, contrary to common perception, wish to spend much more than men on household goods.

While it is true that husbands and wives have relatively similar divisions of expenditures on average, the same does not hold within individual households. In nearly 50% of the couples, the difference between the allocation chosen by the wife and that chosen by her husband is more than 400 Naira. Few couples (around 10%) have an identical allocation. This implies that, while spouses have social preferences, they need not agree on how to divide a budget between specific expenditure categories. Hence, they may wish to influence

¹⁶Thirteen and 20 percent of women allocate less than 1000 to themselves in goods and money respectively and 30 and 22 percent of women allocate less than 1000 to their husband in goods and money respectively. For men, 22 percent and 11 percent allocate less than 1000 to themselves in goods and money and 20 and 36 percent allocate less than 1000 to their wife in good and money respectively.

¹⁷Recall that, although a couple must spend tokens in the stall for which they are designated, we have no way of forcing couples to buy things chosen by either the husband or wife.

household expenditure decisions in the direction of their own social preferences.

4.2 Decision Process

Four variables are used to characterise the decision process of the couple. Two of these (defer and consult) are measured directly; the other two (misrepresent and accommodate) are constructed from the sequence of decisions described above.

The first of these is *deferral* (defer): whether they wish to use final allocation preferred by their spouse. Our findings, summarized in Panel II of Table 3, offer arguably the most striking contrast between spouses in our study: for all split choices, wives are much more likely to defer (67-68%) than husbands (20-26%). The difference is highly significant and there is no noticeable difference in women's proclivity to defer depending on the type of choice. This is consistent with the existence of strong social norms that men's control of household finances is expected to be acknowledged by their wife. Furthermore, women are not less likely to defer even when expenditures are fully rival — i.e., choices A and D — then when they are not – i.e., choices B and C. This suggests that deferral decisions by women are not affected by instrumental considerations, a point we revisit below.

The second measure in our experimental decision tree is the decision to *misrepresent* one's true preferences from the spouse. This is measured as the difference between the communicated split (communicateB-D) and the original split (splitA-D). Panel III of Table 3 presents the average splits that subjects choose to communicate to their spouse. Comparing the allocations reported to their husbands to those they chose when decisions were shrouded, wives allocate slightly more to their husbands in choices B and D and slightly less to themselves in choice C. The magnitude, however, of these changes is small, suggesting minimal misrepresentation. This being said, we also find that, in all three cases, choices communicated by wives shrink allocations towards the middle. A similar, but much more muted, process is observed among husbands.

The third measure is the decision to *consult* (**consult**B-D): subjects are asked whether they wish to be told what allocation their spouse chose. The decision to consult is always secret, but individuals may be planning to accommodate and therefore believe consulting is necessary. Panel IV of Table 3 shows the subjects' willingness to consult their spouse's choice. We see that wives are much more likely to consult than husbands, with few differences across goods. The only surprise is that wives' propensity to consult is less than their propensity to defer — perhaps because consultation is not required when the choice is deferred anyway.

The fourth decision is whether to *accommodate* or dismiss the preferred split communicated by the spouse. Subjects may 'stick to their guns' and keep their original split choice even if it diverges from their spouse's; or they may opt to partially or fully accommodate their spouse's wishes. We regard this measure as the closest to the concept of 'procedural empowerment', meaning that a subject feels entitled to impose his or her consumption allocation preferences on their spouse. To capture this idea, we construct a categorical variable that compares the revised split (**revisedB-D**) to the original split (**splitA-D**). If the subjects 'sticks to their guns', then their revised split is the same as their original split; if they fully accommodate their spouse's choice, their revised split is equal to the communicated split of their spouse. Partial accommodation is when the revised decision goes in the direction of the spouse's choice, but not completely. Over-accommodation is when the subjects are contrarian in the sense that they revise their allocation away from their spouse's preferred choice.

Panel V of Table 3 shows the revised splitting decisions of husbands and wives after having been shown their spouse's communication. We see husbands allocating significantly less than wives to female goods in choice C while in choice D wives allocate more money to themselves than their husbands do. For both women and men, however, these differences were already present in their original splitting decisions.

In Panel VI of Table 3 we examine in more detail the extent to which husbands and wives accommodate the communicated split of their spouse when revising their original choice. Here, accommodation is measured compared to the communication received from the spouse: do subjects move toward their spouse's preferences? We see that mean accommodation by wives is much higher than that of husbands, and the difference is highly significant. This is presented in more detail in Appendix Table A2. We see that full accommodation is the modal behaviour for women: in 57, 60 and 38 percent of choices. In contrast, zero accommodation despite discordant choices is the overwhelming response of husbands (84, 84 and 91 percent of choices).

What these results indicate is a strong procedural inequality between husbands and wives in the study area — but much less inequality in the allocation of consumption expenditures. This suggests that intrahousehold allocative fairness is achieved through social preferences—which are largely (albeit not fully) shared between husbands and wives— not through procedural equality.

4.3 The demand for agency

The discussion so far has focused on decisions where the interests of the two spouses are potentially divergent. In this context, agency has instrumental value because it allows each spouse to allocate the household's consumption budget in a way more in line with their preferences. To investigate the possibility of non-instrumental demand for agency, we examine the choices of food and drink that they make in Block 5. The stated preferences indicate which of the four private consumption bundles c_i , has the highest private utility $U_i(c_i)$ for individual *i*. Since consumption is non-rival — what the husband consumes does not affect the wife's choice set, and vice versa — interfering with the consumption decision of a spouse has no instrumental value. Hence if *i* defers his/her consumption decision to spouse *j*, it can only satisfy *j*'s desire for control. This gives us a clean measure of deferral (deferF) for non-instrumental reasons, i.e., as a way for *i* to increase *j*'s non-material utility from the decision process itself.

In the last line of Panel II of Table 3 we see that 68% of wives and 19% of husbands delegate the selection of their food and drink to their spouse. Do they believe their spouse knows what they want? Since there are only two options, congruent choices should occur with a 50% probability if couples do not know each other's preferences. This is indeed what we find: the proportion of congruent choices is 54, 54, 52 and 49% for husbands and wives for cookies and husbands and wives for drink, respectively.¹⁸ Partners do a bad job of picking for their spouse. This suggests that, on average, deferral has a material utility cost for subjects: they are less likely to consume the items they prefer.

In Table 4, we examine the rate at which subjects defer when they are faced with varying costs of deferral. Since the change in the size of the budget represents a cost of deferral, observed choices map out the demand for agency as a function of the cost.

In domain A (women's versus men's goods) subjects made three sets of deferral choices: 1) when deferral used the spouse's decision over 2500 compared to retention using their own decision over 2100; 2) when the budgets were both 2500 and 3) when deferral used the spouse's decision over 2100 compared to retention using their own decision over 2100. In domain D (money for the woman versus money for the man) subjects made two sets of deferral choices: 1) when deferral used the spouse's decision over 2500 compared to the retention using their own decision over 2100; and 2) when the budgets were both 2500.

In domain J (non-rival choices over food and drink) subjects made one choice but with two randomly allocated deferral budgets: 1) deferral was for two cookies and a full glass of juice while retention was for one cookie and half a glass and 2) deferral and retention had the same budget. Note that, in domain J, no subject faced both choices: the differences are across, not within subjects.

Table 4 shows a downward sloping demand for the choice to defer for women: the higher the price of deferring (smaller budgets for deferring), the less likely women are to defer their

 $^{^{18}}$ In a similar experiment, Afzal *et al.* (2022) show that subjects are no better informed about the preferences of their spouse than a stranger.

choice. In contrast, there is no change for men and most men chose to retain their choices, even when there is a cost for doing so — a low and price-inelastic demand for deferral.

It is possible that women (and men) defer because they believe their spouse will chose for them something that is more advantageous than what they would have chosen for themselves. We investigate this possibility here. We observe both what someone chooses for themselves and what their spouse chooses over the same choice set and we can compare the individual payoffs for both choices to see if deferring is, indeed, sometimes optimal. Since subjects were not asked what they expect their spouse to do in case of deferral, we do not know what each of them expects to gain or lose from deferral but if subjects form rational expectations on splits, on average their guesses should be correct. With these assumptions, we can compare subjects' private material payoff without deferral to the private material payoff they would receive if they deferred. For Table 5 we calculate the proportion of cases in which it would be optimal for subject to defer. We see that, on average, women over-defer relative to what would be optimal, a difference that is always significant. In contrast, husbands massively under-defer: 20-21% when it would be optimal to defer in 60-64% of optimal cases, a difference that is always significant.

Taken together, these results indicate that deferral is driven primarily by non-instrumental considerations. This is particularly clear for men, who seldom defer and, when they do, show no responsiveness to instrumental concerns. Women, in contrast, tend to over-defer. There is no evidence that deferral is a method for improving material outcomes, but, like a normal good, demand for deferral does respond to costs. This serves as further confirmation that the deferral decision is mostly driven by intrinsic motives, such as respect for the husband's authority or adherence to an internalised social norm.¹⁹

4.4 Aggregating agency at the level of couples

So far we have examined individual decisions related to agency. We now turn to the impact on the collective behaviour of spouses that is implied by our experimental results.

Executive agency We start by examining the combined deferral behaviour of spouses across five experimental decisions: A, B, C, D, and J (see Panel II in Table 3). Note that although the proportion of female and male deferrals are quite similar across A, B, C, D, and J (see Panel II in Table 3), they are not perfectly correlated within subject; for A to D,

¹⁹We cannot entirely rule out the possibility that women's adherence to local social norms is reinforced by a kind of Hawthorne effect: being observed encourages participants to be on their best behaviour. We can, however, rule out experimenter effects: neither female nor male participants show any inclination to follow Western gender norms.

correlations coefficients vary between 0.51 and 0.60 for women and between 0.36 and 0.58 for men. Correlation coefficients between deferral in domains A through D and deferral in the food-and-juice experiment are even lower: between 0.22 and 0.31. In addition, deferral decisions are not synchronised across spouses: in only about half of the couples do we observe the wife deferring to her husband and the husband not deferring to his wife, and this is true for all five experimental decisions. In 26-28% of couples, neither the wife nor the husband defer; in 12-14% of them, both husband and wife defer; and in 8-11% of them, the husband defers and the wife does not. Thus, there is significant variation in deferral within individuals and little evidence of coordination in deferral behaviour across spouses.

Consultative agency Next we turn to our measure of consultative agency, C_w^h , which measures how much of the gap $(c_w^w - c_w^h)$ between the preferred allocations of the wife and her husband is bridged by the husband through the process of misrepresentation, consulting, and accommodation. To recall, $C_w^h = 1$ means that the allocation implemented through consultation is that preferred by the wife, and $C_w^h = 0$ means it is that of the husband. A number between 0 and 1 implies a compromise between the two. The husband's consultative agency C_h^w is computed in a similar manner. By construction, C_w^h and C_h^w cannot be computed if $c_w^w = c_w^h$. This happens in around 10% of the cases – see column 1 of Table 6.

In column 2 of Table 6, we report the average values of C_w^h and C_h^w for choices B, C, and D. We see that C_w^h is very low – between 2.6% and 5.2%, indicating that the consultation process reproduced in the lab moves the allocation chosen by the husband only 2.6 to 5.2% in the direction of the wife's preferred choice. In contrast, C_h^w takes values between 31.7% and 40.2%, indicating that wives grant a lot more consultative agency to their husband when given executive agency by the experimenter.

Next we decompose C_w^h and C_h^w in three separate components (see equation 6 which we reproduce below for convenience): misrepresentation (i.e., self-censoring); willingness to consult the spouse; and accommodation when consulting.

$$C_w^h = 1 - \Big(\frac{c_w^w - c_w^{w_comm}}{c_w^w - c_w^h} + \frac{(1 - k_h)(c_w^{w_comm} - c_w^h)}{c_w^w - c_w^h} + \frac{k_h(c_w^{u_comm} - c_w^{h_accom})}{c_w^w - c_w^h}\Big)$$

The results are shown columns 3, 4 and 5 of Table 6. Column 3 reports first term $\frac{c_w^w - c_w^{w.comm}}{c_w^w - c_w^h}$. This term can lie outside the c_w^w to c_w^h range. We see that, for choices B and C – which involve allocating a budget between a private good or a household good – there is little misrepresentation/self-censoring: values are close to 0. Things are different for choice D – which divides a monetary budget between the spouse: here wives reduce what they report to their husband by 18.9% of the difference between their private preferred allocation c_w^w and that of their husband. In contrast, all values are uniformly small for husbands, indicating no self-censoring on their part.

The second term $\frac{(1-k_h)(c_w^{w-comm}-c_w^h)}{c_w^w-c_w^h}$ captures the loss of consultative agency because the spouse refuses to be informed of the participant's chosen allocation.²⁰ It is shown in column 4. We see that this is where most of the loss of consultative agency occurs: for wives, this term accounts for between 67.6 and 83.8% of the difference between $c_w^w - c_w^h$; for husbands it is smaller but still represents between 45.6 and 46.9%.

The third term $\frac{k_h(c_w^{w.comm}-c_w^{h.accom})}{c_w^{w}-c_w^{h}}$ is shown in the fifth column and represents the loss of consultative agency due to lack of accommodation by a consulting spouse. This term is not negligible, accounting for between 10.9 and 18.2% of the gap for wives and between 6.2 and 28.4% for husbands. But it is smaller in magnitude than the second term largely because spouses seldom consult: 15 to 20% of husbands and 52 to 56% of wives consult, depending on the allocation decision.

The experiment allows us to compare C_w^h and C_h^w to the counterfactual allocation $\hat{C}_w^h = 1 - \frac{c_w^u - c_w^{h - accom}}{c_w^w - c_w^h}$ that would result if spouses were forced to hear the allocation that their spouse communicated to them. The results of this calculation are presented in column 6 of the Table. It shows a slight increase in the consultative agency of women, rising by a factor of 2.4 to 3.8 times relative to the values reported in column 2. The modesty of this improvement reflects the fact, already documented earlier in the paper, that husbands do not accommodate the communicated allocation of their wife. In contrast, forcing consultation would improve the consultative agency of husbands by 45 to 68% – assuming counterfactually that they would delegate executive agency to their wife.

Table A3 in the online appendix examines whether material allocations of consumption budgets lean more towards women's preferences when they have executive vs consultative agency – i.e., we compare c_w^w to $k_h c_w^{h_accom} + (1 - k_h) c_w^h$. We find that the material agency of women is slightly but significantly better when they have executive agency, but the differences are quite small on average. Additional analysis, not presented here, nonetheless shows that these small differences in the *average* allocation hide large within-couple differences: how husbands allocate the consumption budget is rarely what wives would have chosen. This means that, for women in our sample, lack of executive agency is consequential: consumption is not allocated the way they want.

 $^{^{20}}$ Note that, because we always show participants the communication of their spouse, this is not observed in the experiment, but can be inferred.

4.5 Allocative efficiency in production

We now turn to Block 6, which allows allocative inefficiency in production, either due to a desire to increase one's individual income, or driven by fairness considerations in input allocation. Most of the theoretical literature on intrahousehold allocation implicitly or explicitly assumes efficiency in production decisions. Yet some evidence suggests that input endowment effects can impede production efficiency of households in Africa (see Udry, 1996).

The two production games played in Block 6 examine whether subjects make decisions that are efficient for the couple. Recall that, in one game, maximising household production requires giving more input to the spouse and in the other, maximising household production requires giving more input to oneself. We see from Panel VII of Table 3 that, in three of the four cases, about half of the subjects choose the efficient outcome that maximises their joint income. The one exception is that, unlike female subjects, male subjects are less likely to choose the efficient allocation when it yields less for themselves — a statistically significant gender difference. This is a priori surprising: since the subjects are paid in cash in front of each other for this block, they could easily compensate each other for making efficient choices after the experiment. The fact that most men do not opt for the efficient outcome when it benefits their wife suggests that they do not, in fact, expect to be able to costlessly appropriate the surplus ex-post.²¹

Combined with our earlier results showing that subjects often defer to their spouse even when doing so reduces the household surplus, these findings indicate the existence of preferences over the process by which a particular allocation is obtained: it is preferable to choose the "right" allocation rather than the "best" outcome and, ex post, reallocate. With this understanding in mind, we now examine whether offering an unconditional cash transfer to women modifies these preferences.

5 Effect of the UCT on intrahousehold agency

Equipped with a better understanding of the decisions made in the lab by couples who did not receive the unconditional cash transfer (UCT) intervention, we now examine the effect of the UCT intervention on female empowerment using the full sample of treated and control participants.

The impact evaluation study by Papineni *et al.* (2024) examines the effects of cash transfers on various economic and social outcomes, with a particular focus on women's empowerment. The findings reveal that, one year after the program ended, beneficiary women

 $^{^{21}}$ In Online Appendix B we examine the full set of choices and show that the best model of decision–making is a blend of preferences over both input and output allocations.

experienced a 20 percentage point increase in enterprise ownership, while non-beneficiary women saw a 13 percentage point increase. This surge in female-led entrepreneurship significantly enhanced household consumption and reduced food insecurity. However, while the cash transfers effectively mobilised previously underutilized female labour, they did not significantly alter societal attitudes toward women working outside the home.

As part of the impact evaluation of the UCT intervention, RCT participants were asked a series of questions about decision-making in their households before and after receiving the program. These questions were used to create indices of empowerment based on the Women's Empowerment in Agriculture Index (WEAI).²² The impact evaluation study by Papineni *et al.* (2024) reports significant positive impacts on this index for the full RCT sample. Specifically, the cash transfer treatment led to a 0.69 standard deviation increase in the pro-WEAI index for beneficiary households compared to the control group.

In Table 7 we reproduce these findings for our much smaller sample of lab participants: while the A-WEAI index constructed from answers to the baseline survey is balanced across UCT treatment and control, we observe a significant effect on pro-WEAI, the index collected at endline – implying an improvement in female empowerment in our sample as well.

Significant improvements are also reported by Papineni *et al.* (2024) for specific components of the pro-WEAI index, such as group membership, inputs into productive decisions, and ownership of assets and plots. Additionally, there was a 0.28 standard deviation increase in the decision–making index, which captures control over enterprise activities and income use. Overall, this shows that the cash transfer program significantly enhanced women's economic empowerment by increasing their participation in decision–making and control over resources.

In spite of these effects, the program did not significantly impact psychosocial components of the index, like self-efficacy and attitudes toward gender-based violence. In this section, we complement these findings by testing whether the UCT intervention affected the gender distribution of executive and consultative agency within couples, separately under the nosecret and secret conditions.

5.1 Testing strategy

To investigate the effect of the UCT treatment and the secret condition in the lab, we estimate a model, in equation 7, that examines the interaction of the secret condition (S) and the UCT treatment (T) as four independent categories (with $T_i = 0$ and $S_i = 0$ as the

 $^{^{22}}$ One, referred to as A-WEAI by Malapit *et al.* (2019), is based on six questions collected at baseline, and another, referred to as pro-WEAI, is based on 12 different questions asked at endline.

omitted category) and also, in equation 8, as the interaction of two treatments.

$$x_i = \beta_0 + \beta_1 [T_i^1 S_i^0] + \beta_2 [T_i^0 S_1^1] + \beta_3 [T_i^1 S_i^1] + \epsilon_i$$
(7)

$$x_i \qquad = \beta'_0 + \beta'_1 T_i + \beta'_2 S_i + \beta'_3 T_i S_i + \epsilon'_i \tag{8}$$

 x_i is any of the budget allocations or other decisions made by the subject, variables T_i and S_i indicate the UCT treatment and secret condition and ϵ_i is an error term which we cluster at the level of the session to control for unobserved session effects that would lead to correlated effects. Note that $\beta_1 = \beta'_1$, $\beta_2 = \beta'_2$, $\epsilon_i = \epsilon'_i$ and $\beta_3 = \beta'_1 + \beta'_2 + \beta'_3$. Since the two regressions are almost exactly the same, we only report the coefficients for equation 7 and the *p*-value of the coefficient β'_3 .

Similar regressions are estimated for the two subsequent allocation decisions subjects are asked to make. The three dichotomous measures of procedural empowerment — defer, consult, and accommodate — are regressed on treatment. If the UCT treatment empowers women in a procedural sense and makes them more openly assertive, we should observe that their decisions to defer, consult, and accommodate all fall with treatment. The need to misrepresent would also fall. Furthermore, if the treatment also reduces the power of husbands, or makes them take their wife's wishes into consideration, we should observe increases in the rate in which men defer, consult, and accommodate.

5.2 Budget allocation decisions

We begin with the initial split decisions of husbands and wives. As explained in Section 3, initial split decisions (Block 1) are always kept secret in the experiment. Thus, choices made by subjects can be interpreted as representing their true preferences and any treatment effect we find can be seen as a shift in preferences induced by the UCT.²³

To investigate this, we show in Table 8 the estimates obtained by regressing split decisions on a UCT treatment dummy. Since choices are shrouded, we expect no effect of the nosecret treatment.²⁴ We see that women who experienced the UCT increased the allocation to themselves in goods and money by 67 Naira in domains A and D, for a total change of 134 Naira. Only the total allocation is significantly different between those who received the UCT and those who did not (the p-values for domain A and D are 0.103 and 0.107 respectively). The total change in A and D for men, by comparison is 12 Naira, an amount

 $^{^{23}}$ A shift in preferences may arise because the utility function of the subject has changed or because of an income effect induced by the UCT.

²⁴We confirm in Online Appendix Table A4 that the secret/no-secret treatment has no separate significant effect on split decisions.

both statistically and economically non-significant. There is also suggestive evidence that women allocated more to the household versus their husband (domain B). The increase in household allocation is 51 Naira (p-value = 0.116)

The UCT treatment changed the amounts women are willing to allocate to themselves (and their household compared to their husband). The changes are marginally significant, but economically small. The total change in the amount women allocate to themselves after a eighteen-month cash transfer program is 134 Naira, 5% of the average allocation across domains A and D. Since men do not change their allocation, the treatment can be seen as widening (slightly) the divide between women's and men's preferred consumption bundles. This finding is consistent with a slight rise of individualism among women who received the UCT.

5.3 Decision Process

In Table 9 we estimate regression model (7) for deferral decisions. Unlike the split decision, defer decisions are only shrouded in the secret condition, something participants were informed of before making a decision. Consequently, we include both a UCT and a secret effect. We find that, in all four rival domains, wives defer much less often when they receive the UCT treatment under the secret condition. The difference is large in magnitude – between 12 and 17 percentage points — and it is significant in all cases and in the aggregate (column A-D). No such changes are observed for husbands who, as we noted earlier, are much less likely to defer on average. We see the results of the same magnitude for the nonrival food and juice (J) but the coefficients are not significant. Combined with the findings from Table 8, these results suggest that the UCT treatment has made women want more consumption and agency. But it has not made them more vocal or openly assertive: they continue to defer to their husband if this decision is observed. This is consistent with deferral being a social norm internalised by participating women, since it is also present in the Secret treatment. The fact that this behaviour changes with the UCT indicates that the treatment changed this internalised norm for some female recipients, but did not change the fear of possible retribution for breaking it. This is quite a remarkable outcome, and not one that would be observable outside this experiment.

Table 10 presents a similar analysis for the decisions to misrepresent, consult, and accommodate. Note that the initial allocation is always secret, the communicated allocation is never secret and the decision to consult is always secret. The decision to accommodation could be revealed in the no-secret condition, since the spouse knows both their communication and the final allocation. Despite the fact that the secret condition only directly affects the accommodation decision, it might indirectly affect misrepresentation and consult, and therefore we examine all four treatment conditions.

In the decision to misrepresent, presented in Panel A, the dependent variable is the difference between the communicated split and the initial (secret) split: a negative coefficient implies that the communicated split is smaller than the secret split. We find that women who received the UCT tend to misrepresent more in the secret condition but the effect is only significant at the 10% level in one of the three regressions. For husbands we observe more misrepresentation as a result of treatment, but only in choice D. The lack of systematic pattern across choices makes us suspect the result is not robust. There is no evidence for an effect of the UCT treatment alone.

In the second panel of Table 10 the dependent variable equals 1 if the subject manifested a desire to see their spouse's communicated split. In five of the six regressions, we observe a large fall in the likelihood of consultation as a result of treatment, with all but one significant effect being concentrated in the secret condition. The magnitude of these effects is large, especially in the secret condition: wives reduce the likelihood of consulting by a combined 13 to 27 percentage points, while husbands reduce it by 16 to 23 percentage points (starting from a much lower base.) The fact that these changes are significant primarily in the secret condition suggests the presence of hidden tension between treated spouses, tensions that they are trying not to learn about.

The third panel of Table 10 focuses on accommodation. Here we find little evidence of treatment effects: except for one significant coefficient at the 10% level, there is no dominant pattern across choices. Even in secret, treated women are not less willing to accommodate their husbands' communicated allocation; the pressure is too direct.

These findings contribute to a coherent picture of the effect of treatment on procedural fairness in couples: treated women become secretly more demanding and less willing to defer and consult, as long as these decisions can be hidden; treated husbands also tend to shift their budget allocation towards a more selfish posture, although the effect is not statistically significant; and they consult less often, especially if this decision is less observable.

5.4 The Demand for Agency

In the control group, we saw that there is a demand for agency: the willingness to defer is sensitive to the cost of deferral. In a parallel analysis, we show that the secret treatment shifts the demand curve, increasing the demand for agency at every price. While changes in consumption bundles or increased demand for agency could reflect an income effect of the UCT treatment on household assets and female labor force participation, the impact of the secret treatment cannot be an income effect as the secret treatment was randomised across individual on the day of the experiment.

We examine the effect of the UCT and no-secret treatments in Table 11 where we combined all the deferral decisions subjects make in domains A (female v. male goods) and D(cash for the wife v. the husband) and in the food and drink domain J. By combining decisions made in Blocks 2 and 3, we can estimate the response of the demand for agency (i.e., non-deferral) to its cost.²⁵ The results confirm what we found earlier: that increasing the cost of agency reduces women's willingness to exert agency. For men, the cost of agency has no significant effect on deferral, except for domain D (cash to the wife) where increasing the cost to deferral actually *increases* the propensity to defer.²⁶

In the secret condition the deferral decision cannot be discovered by the spouse; otherwise, it can. We expect the secret condition to potentially have an effect on deferral: if a wife is enticed to demand more agency as a result of the UCT treatment but is afraid of retribution if it is revealed that she did not defer to her husband, she may refrain from deferring only in the secret condition. The impacts of the treatment match those shown above: women (but not men) increase their demand for agency if they received the UCT treatment and are in the secret condition. This result is significant both in total — i.e., UCT recipients in the secret condition are different from non-recipients in the no-secret condition — as well as on the margin — i.e., the combined impact of the UCT and secret treatments is different from the sum of the UCT treatment and secret condition.

5.5 The distribution of agency within the couple

So far we have examined the impact of the UCT and secret treatments on individual behaviour. We now turn to its impact on the collective agency of spouses. We start by creating four variables that average the proportion of paired spousal decisions across experimental choices A to J. The first variable represents the proportion of paired decisions in which the wife defers and the husband does not, and so on. Equal agency is achieved when both defer – and it can be seen as a premise for a negotiated outcome. In contrast, disharmony in the couple can be expected when neither defer. We already know that, on average, the combination of UCT and secrecy induces less deferral by wives without changing husbands' behaviour. But these averages could mask shifts in opposite directions across couples: if men who deferred to a non-deferring wife no longer do as a result of treatment (i.e., they

 $^{^{25}}$ Coefficient estimates for the cost variable in the *J* domain are across-subject comparisons since no subject made a decision over deferring a cookie and drink at two different costs.

²⁶This can be read as a reaction to experimental demands for deferral, but should not be interpreted as a reaction to female agency; men are pushing back on the experimenter, not necessarily their wives.

oppose the treatment), we could observe both an increase in the proportion of equal agency couples (both defer) as well as in the proportion of disharmonious couples (neither defer).

To investigate this issue, we regress each of these four proportions on the UCT and secret treatments $.^{27}$ Results, shown in Table 12, show that the combination of the UCT and the secret treatment causes a transfer of around 8-9% of couples from the "Wife defers, husband does not" to "No one defers" – and another 4-5% transfer from "Both defer" to "Husband defers, wife does not" (significant only for the former). These findings indicate no correlation between deferral among husbands and the change in deferral by their wife – i.e., it is not the case that husbands react negatively to a reduction in deferral by their spouse. We nonetheless note a reduction in the proportion of couples who both defer from around 16% to 11% and an increase in the proportion of no-deferral couples from approximately 23% to 32%. In other words, in our sample, equal agency has gone down and the risk of conflict has increased – a finding that is largely due to the fact that husbands do not accommodate the increased demand (or rising aspirations) for women's agency that the UCT induced. Given this lack of accommodation, it is no surprise that pent-up demand for executive agency remains hidden: it is only expressed in secret.

May husband nonetheless have accommodated their wives' increased demand for agency by consulting them more? To investigate this possibility, we regress C_w^h and C_h^w on the UCT and secret treatments, as done in Table 12 for deferral. We find no evidence on an effect of treatment on the consultative agency of either husbands or wives.²⁸ The same holds for the aggregate consultative index when the spouse is forced to consult. We also find no evidence of any effect of the UCT or secret treatments on material allocations resulting from consultation among spouses.²⁹ This indicates that the UCT intervention has had no effect on husbands' willingness to consult their wife over the allocation of consumption in the household, suggesting that the intervention has not changed gender norms in favor of women. This evidence also contradicts the collective thinking hypothesis: if collective thinking was the reason for women not to want to exert executive agency over consumption, then we should observe husbands consulting them – which we do not.

5.6 Aggregate Welfare and Secrecy

Here, we examine the effect of treatment on the expected payoff of experimental subjects. As explained in Section 2, one of a long list of possible choices made by subjects is drawn at

²⁷Similar but slightly more significant results obtain using a fractional logit model.

 $^{^{28} {\}rm Of}$ the 18 estimated coefficients (3 regressors x 3 decisions x 2 spouses), only one is significant at the 10% level – see Appendix Table A5.

²⁹See Appendix Table A6.

random to determine final payoffs. We focus on the most important choices made by both spouses, namely **split**A-D and **defer**A-D. Using simple rules described in the note to Table 13, we construct a variable that measures the payoffs π_f and π_m that a male and female subject can expect to receive based on their own decisions. These expected payoffs are what subjects can expect to receive purely for themselves at the end of the experiment based on their own split and defer decisions and the decisions of their spouse.

Coefficient estimates are presented in Table 13. Although the point estimate for the UCT is positive, it is not significant. The material payoff of women only increased in the UCT/secret treatment. The increase in this category is about 5 percent of the total possible payouts. To the extent that the rule for selecting payoffs in the experiment mimics what happens at home, this suggests that the UCT treatment is most likely to raise the material welfare of women if it is accompanied by some form of secrecy. Without secrecy, women overwhelmingly delegate consumption decisions to their husband, whose social preferences are largely unaffected by treatment.

5.7 Allocative Efficiency in Production

Finally, we examine the impact of the UCT treatment and other experimental conditions on the efficiency of couple's production decisions. Based on the existing literature, we have no reason to believe that the UCT treatment should increase or decrease efficiency, since theory suggests efficiency is a dominant strategy in all households, irrespective of the decision– making process. Moreover, the game does not allow deferring production decisions to one's spouse, so there is no way to express or avoid agency — and thus no possible effect of the UCT through demand for agency.

Table 14 summarizes the impact of the UCT treatment and secret conditions on decision– making in the production games. We examine two measures of efficiency: subjects who played both versions of the game are strongly efficient if they chose the efficient outcome in both games; subjects who played only one version are weakly efficient if they chose the efficient outcome in that game.

In the top Panel of Table 14, we see that, among control participants (no-UCT) in the secret condition, women chose the efficient option less frequently. This is true for both sets of subjects, i.e., those who played both versions of the game and those who only played one version. Based on model predictions, this finding indicates lower social preferences for women in the condition that mimics a non-cooperative state. There is no similar effect of the secret condition on men (see Panel II) but, on average, they choose the efficient option less frequently — suggesting that, unlike women, men act equally selfishly in both the secret

and no-secret conditions.

Turning to the UCT, we see that it induces a large drop in the frequency of efficient choices by women in the no-secret condition. Since this loss in efficiency is potentially visible to the husband, it suggests that, on average, treated women take a more confrontational stance in the production game. The effect of the UCT is also negative in the secret condition, but much less so than among controls, and it is no longer significant. This, in contrast, seems to suggest that women who received the UCT become more altruistic in secret. We similarly find that husbands of UCT recipients choose the efficient option more often in the secret condition (although not significantly so) but we do not see this in the no-secret treatment. This implies that, like women, men in UCT households behave more altruistically in the secret than in the no-secret condition. When combined with the previous result that consulation fell in the UCT, these results suggest that the UCT treatment induced a fall in cooperative outlook — but an increase in altruism — among both spouses.

6 Conclusion

The use of an experimental laboratory setting allowed us to examine changes in decision– making processes after a year-and-half-long unconditional cash transfer, to break decisions into constituent parts and to test how decisions would be made both in a secret and no secret condition. Although we acknowledge the possible limitations of laboratory evidence in terms of external validity, the approach allows us to distinguish between key components of the intrahousehold decision process, namely, the social preferences of the spouses and the distribution of executive and consultative agency between them.

By examining the behaviour of participants in the control group we show that both men and women have social preferences over each others' consumption, even with fully rival goods, and we do not see the gender imbalance in outcomes that we had expected in this setting. This stands in contrast with much of the literature on female empowerment which has bypassed the issue entirely by focusing instead on the private material welfare of the spouses. Our findings demonstrate that ignoring the possibility that spouses have social preferences and may behave altruistically towards each other is likely to lead to an incorrect assessment of female empowerment (see also Almås *et al.*, 2020). Indeed, we also find that the *way* allocations are chosen is highly biased. Women are more likely than men to consult their spouse when offered the option, to misrepresent the choices they communicate to their spouse (stating they want less than they originally chose), and to accommodate their spouse's choices after being informed of them. In addition to observing the negotiation process over outcomes, we allowed both men and women to skip this process by deferring their decision to their spouse. We find a marked contrast between men and women. Women defer 67% of their decisions to their spouse while men only defer between 20 and 25% of their decisions. Furthermore, men refuse to defer even when the allocation is a choice between household items and female items, a choice over which they likely have little at stake in the context of our experiment.

Our experiment demonstrates that intrahousehold equity in consumption is achieved through social preferences, not through participation in household decisions. The process by which decisions are made leads to final allocations that are not, on average, different from the initial choices of either spouses. In fact, any examination of the initial and final choices of the couples would give the impression of significant female bargaining power. By varying the costs of decisions and examining their efficiency, both at the individual level and household level, we see strong willingness to incur costs to follow the process outlined above: women over- and men under-defer relative to optimal. This strongly suggests an intrinsic value to the process: in the local context of our experiment, men 'need' to be decision makers and women 'need' to let their husband decide. Men in particular, show no sensitivity to the costs or benefits of deferral.

This leads to inefficiency. Although husbands on average make choices that are similar to the choices of their wives, women defer to them even when doing so reduces their joint surplus. Husbands are making similarly expensive decisions by failing to defer even when doing so reduces the couple's joint surplus and their wives would have made the same decisions that they made. The inefficiency is apparent in these costly decisions as well as in the two production games. Women are trying to balance inputs and outputs across the couple instead of just picking the clearly largest pie and then reallocating after the fact. Men are also making inefficient decisions and their poor choices are quite visible in the production game where efficiency requires them to let their wives earn more than they do. These men should be able to capture the gain in income but they are unwilling to allow this to happen — a result that is reminiscent of the findings of Bertrand *et al.* (2015) in a developed country context.

We do see some changes in these patterns in households that received the UCT. There are small changes in the initial choices of allocations for both men and women, suggesting some impact on the weights each member places on their own consumption as a result of wives receiving the UCT: in the lab, women tend to allocate a slightly larger share of the consumption budget to themselves, suggesting a change in social preferences (or an income effect). The changes in processes are muted: UCT recipients are less likely to defer their budget allocation decisions to their husband, but only when doing so is not observable. This is compatible with a pent-up demand for agency that would be repressed if expressed openly. We find no effect of the UCT intervention on husbands' propensity to consult their wife or accommodate their preferences, indicating no effect of the UCT on gender norms of power in the household. Taken together, these results suggest that the effect of the UCT on the intrahousehold distribution of agency is limited to a rise in aspiration for agency that is not met by changes in gender norms or fear of reprisal, and therefore can only be expressed in secret – a finding reminiscent of Ashraf *et al.* (2014).

Our results call into question the ability to infer bargaining power by focusing on outcomes: in settings with strong social preferences, outcomes may bear little resemblance to the distribution of power within the household. The UCT has been shown to increase the material welfare of women, children, and the household. But it only had a small impact on the way decisions are made in the household, i.e., that women demand more agency only when they can do so in secret. While not unimportant (many households decisions are shrouded and therefore, effectively secret), this small change does not support the hypothesis that a cash transfer significantly improve agency in the household.

The objective of this paper was to investigate whether participants to our study have a demand for agency, not *why* they do not demand agency. There are many reasons other than social norms why individuals may prefer to delegate decisions to others: because they do not care enough to waste time or cognitive load on making a decision (e.g., Dhar, 1997; Shafir *et al.*, 1993); because others are better informed and thus better able to make a good decision (e.g., Costa-Gomes *et al.*, 2022; Tagat *et al.*, 2023); because they lack self-confidence or have a failure of aspirations; or because they altruistically opt to let others decide who value making decisions. More research is needed to disentangle how these various factors influence decision–making within households.

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| Decision | Experimental flow for one block (B) with tw Description | Ex 1 | Ex 2 |
|------------------------------|--|--------------------------------------|-----------------------------|
| | • | | |
| ${f split}{f B}$ | Block 1 Divide 2500 Naira between male items and household items | 1400/1100 | 1300/1200 |
| $\mathbf{resplit}\mathrm{B}$ | Divide 2100 Naira between male items and household items | 1100/1000 | 900/1200 |
| | Block 2‡ | | |
| $\mathbf{defer}\mathbf{B}$ | Keep original split or change to spouse's | defer | not defer |
| | Block 3‡ | | |
| defercostB | Keep original split or change to spouse's (lower total) resplit | defer | not defer |
| $deferbenefitB^{\dagger}$ | Keep original (lower total) resplit or change to spouse's split | defer | defer |
| | Block 4 | | |
| communicateB | choose allocation to show to spouse | 1200/1300 | 1300/1200 |
| $\mathbf{consult}\mathrm{B}$ | Ask to see the decision of spouse | yes | no |
| shown | All participants shown communication of | (1800/700) | 1200/1300) |
| reviseB | their spouse, whether consult or not Revise split after being shown decision of spouse (even if did not consult) | 1500/1000 | 1300/1200 |
| | Block 5 (Always implemented immedia | telv) | |
| ${\bf split} F, J$ | Choose round or square cookies and coke or fanta for self | round:coke | round:fanta |
| spousesplitF,J | Choose round or square cookies and coke | round:coke | square:coke |
| | or fanta for spouse | | |
| deferF | Defer both cookies and drink to spouse's choice or keep original choice | defer | not defer |
| randomised Payout | One spouse is randomly selected | wife | husband |
| | One decision (split , resplit , defer , de - fercost , deferbenefit) from all four do- mains is randomly selected and imple- | wife's $deferB \downarrow$ husband's | husband's split B |
| Shopping | mented for that spouse Couples receive experimental money to shop, specific to each store | split B 1800/700§ | 1200/1300§ |

Table 1: Experimental flow for one block (B) with two female examples

Notes: Only domain B (male items vs. household items) is shown; ther domains follow the same protocol. † Defer with benefit was not implemented for domain B, shown here for completeness. ‡ The order of Block 2 and Block 3 is randomly assigned within session. § Spouse choices are shrouded by placed individuals in separate rooms and allowing for the possibility that the selection would be replaced with a random decision for any decision made in Block 1 and for all deferral and consulting decisions included in the secret treatment.

| Choice | 1 | Spouse input | | Spouse Income | Total Income | | | | | |
|---------------|---|------------------|-----------------|---------------|--------------|--|--|--|--|--|
| | Game 1: efficient means less for self than spouse | | | | | | | | | |
| a | 10 | 0 | 1500 | 600 | 2100 | | | | | |
| b (dominated) | 7 | 3 | 1200 | 800 | 2000 | | | | | |
| c (efficient) | 3 | 7 | 1000 | 1500 | 2500 | | | | | |
| d | 0 | 10 | 400 | 1900 | 2300 | | | | | |
| | Game 2 | : efficient mean | more for self t | han spouse | | | | | | |
| a | 10 | 0 | 1900 | 400 | 2300 | | | | | |
| b (efficient) | 7 | 3 | 1500 | 1000 | 2500 | | | | | |
| c (dominated) | 3 | 7 | 800 | 1200 | 2000 | | | | | |
| d | 0 | 10 | 600 | 1500 | 2100 | | | | | |

 Table 2: The Production Games

| | | ple mean | | | 1 |
|---|----------|-------------|----------|-------------------------|-----|
| | Wife | Husband | t-stat | p-value | Ν |
| Panel I: Split the budget secretly - spendi | ng on | 1st good | | | |
| ${f split}A-{f female}$ vs male goods | 1410 | 1298 | 3.22 | 0.001 | 251 |
| ${f split}B-$ household vs male goods | 1396 | 1316 | 2.36 | 0.019 | 251 |
| \mathbf{split} C– household vs female goods | 1168 | 1214 | -1.37 | 0.171 | 250 |
| \mathbf{split} D– money for wife vs husband | 1265 | 1062 | 5.67 | 0.000 | 251 |
| Panel II: Decision to delegate splitting de | cision | to spouse: | | | |
| defer A– female vs male goods | 67% | 25% | 10.19 | 0.000 | 223 |
| defer B– household vs male goods | 68% | 22% | 11.22 | 0.000 | 251 |
| defer C– household vs female goods | 67% | 20% | 12.38 | 0.000 | 251 |
| defer D– money for wife vs husband | 67% | 26% | 9.77 | 0.000 | 251 |
| deferF–choice of cookie and juice | 68% | 19% | 10.32 | 0.000 | 165 |
| Panel III: Split shown to spouse - spending | g on 1 | st good | | | |
| communicateB – household vs male goods | 1316 | 1328 | -0.30 | 0.761 | 198 |
| communicateC – household vs female goods | 1178 | 1240 | -1.75 | 0.082 | 223 |
| communicateD – money for wife vs husband | 1226 | 1076 | 2.90 | 0.004 | 123 |
| Panel IV: Whether decides to see the spli | t show | n by spou | lse | | |
| consultB – household vs male goods | 55% | 23% | 6.83 | 0.000 | 198 |
| consultC – household vs female goods | 57% | 20% | 8.63 | 0.000 | 223 |
| consultD – money for wife vs husband | 59% | 20% | 7.24 | 0.000 | 123 |
| Panel V: Revised split after seeing spouse | 's split | : - spendir | ng on 1 | st good | |
| reviseB – household vs male goods | 1329 | 1323 | 0.20 | 0.840 | 198 |
| reviseC – household vs female goods | 1161 | 1230 | -2.70 | 0.007 | 223 |
| reviseD – money for wife vs husband | 1172 | 1089 | 1.72 | 0.088 | 123 |
| Panel VI: Extent of accommodation of spe | ouse's | communio | cated sp | plit^* | |
| accommodateB – household vs male goods | 1.94 | 0.47 | 11.70 | 0.000 | 197 |
| accommodateC – household vs female goods | 1.91 | 0.41 | 12.69 | 0.000 | 223 |
| accommodateD – money for wife vs husband | 1.27 | 0.22 | 6.00 | 0.000 | 121 |
| Panel VII: Whether chooses efficient alloc | ation i | in product | tion | | |
| efficientG1 – efficient is less for self | 56% | 38% | 3.74 | 0.000 | 227 |
| efficientG2 – efficient is more for self | 51% | 53% | -0.48 | 0.633 | 215 |

Table 3: T-tests of the difference between wife and husband means in the control sample

Notes: Each row reports the t-test between sample means, using observations on control households. Observations vary because some decisions were introduced in later sessions. Split choices A to D = split 2500 between two goods. A = female v male goods; B = household v male goods; C = household v female goods; D = money for wife v husband. All goods are purchased in a shop set up in the lab. Resplit decisions are made on a smaller amount (either 2100 or 1800). The choices for communicate (Panel III) and revise (Panel V) are made over the same allocations as those in Panel A. In Panels I, III, and V, the averages shown are the budget amount allocated to the *first* option – e.g., female goods for splitA.

(*) Variable defined as: -1 (move away from spouse's choice); 0 (keep same choice, different from spouse's choice); 1 (keep same choice=communicated choice of spouse); 2 (partially accommodate the spouse's choice); 3 (fully accommodate spouse's choice); 4 (over-accommodate spouse's choice).

| | Rate of deferral | | | | | | |
|---------|------------------|--------|-------------|--------------------------|-----|--|--|
| | | The bu | dget when a | deferring is | | | |
| | Choice: | larger | the same | $\operatorname{smaller}$ | Ν | | |
| Wife | А | 74.1% | 68.3% | 53.2% | 139 | | |
| | D | 73.2% | 71.1% | | 97 | | |
| | J | 93.0% | | | 86 | | |
| | J | | 68.5% | | 165 | | |
| Husband | А | 20.1% | 26.6% | 22.3% | 139 | | |
| | D | 19.6% | 24.7% | | 97 | | |
| | J | 25.6% | | | 86 | | |
| | J | | 19.4% | | 165 | | |

Table 4: The demand for agency (choosing not to defer)

Note: The dependent variable is the choice to defer the allocation to the spouse. The number of observations differs because some treatments only appear in certain sessions. Drink and Juice choices were only posed once for each participant.

| Comparing deferral propensity to optimal own deferral | | | | | | | | | | | |
|---|---------|--------|---------|--------|---------|-----|--|--|--|--|--|
| Domain: retain v. defer | decider | Actual | Optimal | t-stat | p-value | N | | | | | |
| A 2500: 2500 | Wife | 67% | 46% | 4.85 | 0.000 | 223 | | | | | |
| | Husband | 25% | 46% | -5.08 | 0.000 | 223 | | | | | |
| D | Wife | 67% | 44% | 5.46 | 0.000 | 251 | | | | | |
| | Husband | 26% | 44% | -4.65 | 0.000 | 251 | | | | | |
| | | | | | | | | | | | |
| A 2100: 2500 | Wife | 74% | 61% | 2.54 | 0.012 | 153 | | | | | |
| | Husband | 21% | 64% | -8.69 | 0.000 | 154 | | | | | |
| D | Wife | 73% | 49% | 3.47 | 0.001 | 97 | | | | | |
| | Husband | 20% | 60% | -6.35 | 0.000 | 97 | | | | | |
| | | | | | | | | | | | |
| A 2500: 2100 | Wife | 53% | 34% | 3.27 | 0.001 | 154 | | | | | |
| | Husband | 21% | 36% | -2.97 | 0.003 | 154 | | | | | |

Table 5: Examining deferral efficiency

Notes: Deferral choices refer to situation in which the subject must choose between their own split or the split choice made by their spouse. Data uses only observations on control households. Deferral is defined to be optimal if the subject would receive more (or the same) for self by deferring to spouse.

| | Consultative agency | index if spouse | forced to consult | (9) | e Husband | % 63.8% | | % 65.1% | | 6 53.5% | | | lat from the |
|------------------------------|-------------------------------------|----------------------|-------------------|-----|---------------------|-----------|------------------|-----------|--------------------|----------|-------------|---------|---|
| | Con | inde | force | | Wif | 12.5% | | 11.8% | | 6.3% | | | somewh |
| | | Limited accomm. | use | (5) | Wife Husband Wife | 6.2% | | 13.7% | | 28.4% | | | ments differs a |
| | due to: | Limite | by spouse | | | 12.1% | | 18.2% | | 10.9% | | | ee compo |
| | Loss of consultative agency due to: | No consultation | use | (4) | Wife Husband | 46.9% | | 45.6% | | 46.9% | | | um of the thr |
| gency | consulta | No con | by spouse | | Wife | 75.6% | | 83.8% | | 67.6% | | | why the s. |
| Table 6: Consultative agency | Loss of e | isored | comm. to spouse | (3) | Wife Husband | 3.0% | | 3.7% | | -1.0% | | | und 1. This explains w aggregate index value. |
| le 6: Co | | Self-censored | comm. | | Wife | 5.5% | | -4.2% | | 18.9% | | | and 1. T aggrega |
| Tab | Consultative agency | index if spouse free | to consult or not | (2) | Husband | 40.2% | | 40.2% | | 31.7% | | | remain between 0 |
| | Consul | index i | to cons | | Wife | 5.2% | | 3.1% | | 2.6% | | | orized to |
| | % of cases | when spouses' | chosen alloc. | (1) | are identical | 10.1% | | 8.0% | | 10.5% | | | Note: The aggregate index is winsorized to remain between 0 and 1. This explains why the sum of the three components differs somewhat from the aggregate index value. |
| | | | | | Decision: | Household | vs male items | Household | vs temale items | Cash for | wife vs for | husband | Note: The aggr |

| agen |
|--------------|
| Consultative |
| Table 6: |
| |

| | | | | 1 | |
|--|---------|------|--------|---------|-----|
| | | UCT | Treatm | ent | |
| | Control | UCT | t-stat | p-value | Ν |
| A-WEAI index (6 questions at baseline) | 1.94 | 2.09 | -1.42 | 0.155 | 503 |
| pro-WEAI index (12 questions at endline) | 4.42 | 5.02 | -3.96 | 0.000 | 446 |

Table 7: Female empowerment index in the lab sample

Notes: The A-WEIA index stands for Adapted Woman's Empowerment in Agriculture Index. It is based on 6 questions asked to female participants at baseline. The pro-WEAI index stands for Project Woman's Empowerment in Agriculture Index and is based on 12 different questions asked to participants at endline as described in Malapit *et al.* (2019). Both indices capture empowerment but the component indicators in the baseline and endline index are slightly different. The Table only uses respondents to the baseline and endline surveys that also participated to the lab experiment. The A-WEAI index should be balanced across treated and control, while the pro-WEAI index may be affected by the UCT treatment.

| Dependent Variable is Split Decision by the Wife | | | | | | | | |
|--|--------------|---------------|------------|----------|---------------|--|--|--|
| for choice: | A | B | C C | D | A + D | | | |
| UCT | 66.83 | 51.40 | -16.81 | 66.81 | 133.6** | | | |
| | (1.674) | (1.610) | (-0.535) | (1.650) | (2.164) | | | |
| Constant | 1,410*** | $1,396^{***}$ | 1,168*** | 1,265*** | 2,674*** | | | |
| | (61.07) | (49.22) | (40.89) | (41.66) | (74.57) | | | |
| Observations | 503 | 503 | 502 | 503 | 503 | | | |
| R-squared | 0.008 | 0.004 | 0.000 | 0.007 | 0.014 | | | |
| Dependent Varia | ble is Split | Decision b | y the Husb | and | | | | |
| for choice: | А | В | С | D | A + D | | | |
| UCT | -8.921 | 5.295 | -11.96 | -3.422 | -12.34 | | | |
| | (-0.287) | (0.131) | (-0.327) | (-0.104) | (-0.248) | | | |
| Constant | 1,298*** | 1,316*** | 1,214*** | 1,062*** | $2,360^{***}$ | | | |
| | (50.96) | (46.06) | (37.08) | (40.28) | (64.73) | | | |
| Observations | 503 | 503 | 503 | 503 | 503 | | | |
| R-squared | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | | | |

Table 8: Effect of UCT on split decisions

Notes: The dependent variable is the initial split decision (between 0 and 2500) made by each subject in each of the four domains: A: female v. male goods; B: household v. female goods; C: household v. male goods; D: female v. male money. UCT refers to the subjects who received the UCT treatment. Robust t-statistics in parentheses, clustered by session. *** p<0.01, ** p<0.05, * p<0.1

| Table 9: Treatment effects on deferral decisions | | | | | | | | |
|--|---------------|---------------|---------------|----------|---------------|----------|--|--|
| Dependent Variable is Defe | rral Decis | sion by th | ne Wife | | | | | |
| for choice: | А | В | С | D | J | A-D | | |
| UCT/no-Secret | -0.050 | 0.011 | -0.064 | -0.013 | -0.089 | -0.141 | | |
| , | (-0.871) | (0.149) | (-1.112) | (-0.182) | (-1.167) | (-0.619) | | |
| no UCT/Secret | 0.000 | 0.036 | -0.029 | -0.013 | -0.051 | 0.013 | | |
| | (0.006) | (0.486) | (-0.361) | (-0.190) | (-0.578) | (0.046) | | |
| UCT/Secret | -0.143** | -0.123* | -0.124* | -0.168** | -0.133 | -0.591** | | |
| | (-2.148) | (/ | () | (/ | (-1.416) | · / | | |
| Constant | 0.672*** | 0.664^{***} | 0.679^{***} | 0.679*** | 0.709^{***} | 2.707*** | | |
| | (14.561) | (13.710) | (13.096) | (14.424) | (10.563) | (14.571) | | |
| Observations | 450 | 503 | 503 | 503 | 337 | 450 | | |
| R-squared | 0.015 | 0.017 | | 0.022 | | 0.026 | | |
| p-value of marginal effect‡ | [0.280] | [0.075] | [0.729] | [0.098] | [0.952] | [0.126] | | |
| Dependent Variable is Defe | rral Decis | sion by th | e Husbar | ıd | | | | |
| for choice: | А | B | | D | J | A-D | | |
| UCT/no-Secret | -0.052 | -0.030 | 0.008 | -0.092 | -0.049 | -0.196 | | |
| , | (-1.229) | (-0.656) | (0.173) | (-1.578) | (-0.851) | (-1.450) | | |
| no UCT/Secret | -0.007 | -0.037 | -0.023 | -0.097* | -0.129 | -0.203 | | |
| | (-0.104) | (-0.873) | (-0.412) | (-1.858) | (-1.703) | (-1.073) | | |
| UCT/Secret | -0.002 | -0.022 | 0.061 | -0.068 | -0.009 | -0.043 | | |
| | (-0.036) | (-0.484) | (1.040) | (-1.331) | (-0.124) | (-0.235) | | |
| Constant | 0.250^{***} | 0.237^{***} | 0.206^{***} | 0.305*** | 0.256^{***} | 1.034*** | | |
| | (6.783) | (7.842) | (5.181) | (8.142) | (4.426) | (8.495) | | |
| Observations | 450 | 502 | 503 | 503 | 337 | 449 | | |
| R-squared | 0.002 | 0.001 | 0.006 | 0.008 | 0.015 | 0.004 | | |
| p-value of marginal effect‡ | [0.434] | [0.468] | [0.287] | [0.106] | [0.026] | [0.119] | | |

Notes: The dependent variable is the deferral decision made by each subject in each of the four deferA-D decisions and in the deferJ decision from Block 5. A: female v. male goods; B: household v. female goods; C: household v. male goods; D: female v. male money, J: food and drink. The dependent variable is a dummy equal to 1 if the subject chooses to defer, 0 otherwise. The dependent variable in the last column is the sum of the dependent variables in columns A to D; it takes values from 0 to 4.: Each regressor corresponds to a different treatment category with the intercept corresponding to the no-UCT/no-secret category.

 \ddagger is the p-value of the F-test that UCT / Secret = UCT + Secret and tests the marginal contribution of the combined treatments compared to the sum of the contributions.

Robust t-statistics in parentheses, clustered by session. *** p<0.01, ** p<0.05, * p<0.1

| decision by spouse: | | Wife | 1 | , | Husband | |
|--------------------------------------|---|--------------|--------------------------|----------------|---|---|
| for choice: | В | С | D | В | С | D |
| Panel A: Dependent | variable | is Misrep | resent: co | mmunicate | e minus sp | litB-D |
| UCT/no-Secret | -24.44 | -8.692 | 7.750 | 8.095 | 2.647 | -23.25* |
| | (-0.712) | (-0.339) | (0.163) | (0.666) | (0.170) | (-1.814) |
| no UCT/Secret | -42.58 | -7.740 | 17.06 | 40.68 | -33.54 | -9.524 |
| | (-1.309) | (-0.247) | (0.293) | (1.692) | (-1.460) | (-0.561) |
| UCT/Secret | -20.63 | -50.47^{*} | -43.62 | 2.058 | 0.0556 | -26.43** |
| | (-0.444) | (-1.826) | (-1.078) | (0.118) | (0.004) | (-2.192) |
| Constant | -13.33 | 14.41 | -25.40 | -18.10 | -1.695 | 9.524 |
| | (-0.738) | (0.697) | (-0.748) | (-1.624) | (-0.113) | (1.261) |
| Observations | 394 | 450 | 245 | 394 | 450 | 245 |
| p-value of marg. effect [‡] | [0.473] | [0.418] | [0.389] | [0.097] | [0.217] | [0.800] |
| Panel B: Dependent | variable | is Consul | t | | 2 2 | |
| UCT/no-Secret | -0.0365 | -0.0514 | -0.0626 | -0.0556 | -0.0468 | -0.163** |
| 1 | (-0.472) | (-0.663) | (-0.729) | (-1.483) | (-1.035) | (-2.272) |
| no UCT/Secret | -0.0648 | -0.0419 | -0.117* | -0.0839 | -0.0849 | -0.0556 |
| , | (-0.848) | (-0.575) | (-1.776) | (-1.267) | (-1.265) | (-0.688) |
| UCT/Secret | -0.0904 | -0.150** | -0.172*** | -0.125** | -0.114** | -0.0673 |
| 1 | (-1.412) | (-2.150) | (-3.204) | (-2.132) | (-2.461) | (-0.838) |
| Intercept | 0.581*** | 0.585*** | 0.651*** | 0.267*** | 0.237*** | 0.222*** |
| - | (12.10) | (11.87) | (27.14) | (6.924) | (5.597) | (3.619) |
| Observations | 394 | 450 | 245 | 394 | 450 | 245 |
| p-value of marg. effect [‡] | [0.915] | [0.573] | [0.944] | [0.838] | [0.794] | [0.105] |
| Panel C: Dependent | variable | is Accom | modate se | e Table 3 | | L _ |
| UCT/no-Secret | 0.0315 | 0.0162 | -0.0252 | -0.00794 | -0.0616 | 0.0177 |
| | (0.482) | (0.355) | (-0.271) | (-0.123) | (-1.059) | (0.285) |
| no UCT/Secret | -0.0409 | 0.00670 | -0.0452 | -0.0363 | -0.00444 | -0.0421 |
| , | (-0.587) | (0.092) | (-0.631) | (-0.615) | (-0.0749) | (-0.823) |
| UCT/Secret | -0.0406 | -0.0662 | -0.0114 | -0.0209 | -0.0474 | -0.0883* |
| , | (-0.551) | (-1.171) | (-0.139) | (-0.341) | (-0.784) | (-1.781) |
| Intercept | | 0.219*** | 0.669*** | | 0.476*** | 0.159*** |
| - | (13.25) | (4.456) | | (4.397) | (8.594) | (3.985) |
| Observations | 392 | 449 | 243 | 394 | 450 | 245 |
| | | | | [0.776] | | |
| - | $\begin{array}{c} 0.654^{***} \\ (13.25) \end{array}$ | (4.456) | 0.669^{***} (17.01) | (4.397) 394 | $\begin{array}{c} 0.476^{***} \\ (8.594) \end{array}$ | $\begin{array}{c} 0.159^{***} \\ (3.985) \end{array}$ |

Table 10: Treatment effects on decisions to misrepresent, consult, and accommodate

Notes: Each regressor corresponds to a different treatment category with the intercept corresponding to the no-UCT/no-secret category. As explained in Section 3, subjects are first asked whether they want to consult their spouse allocation and this decision is recorded. Later they are told what their spouse chose and are offered an opportunity to revise their chosen allocation. This design allows us to observe accommodation for all subjects. \ddagger is the p-value of the F-test that UCT/Secret = UCT + Secret and tests the marginal contribution of the combined treatments compared to the sum of the contributions. Robust t-statistics in parentheses, clustered by session. *** p<0.01, ** p<0.05, * p<0.1

| Dependent variable is deferral by | wife | husband | | | |
|---|---------------|---------------|--|--|--|
| A (female v. male goods) | omitted | | | | |
| D (cash for wife v. husband) | 0.0161 | 0.0172 | | | |
| | (0.776) | (0.874) | | | |
| J (drink and cookie) | 0.0273 | -0.0140 | | | |
| | (0.951) | (| | | |
| A X cost | -0.0220*** | | | | |
| | (-6.009) | (-0.491) | | | |
| D X cost | -0.0150* | 0.0129^{*} | | | |
| | (-1.773) | (1.813) | | | |
| J X cost | -0.0623*** | -0.00106 | | | |
| | (-6.989) | (-0.109) | | | |
| UCT/no-Secret | -0.0182 | -0.0417 | | | |
| | (-0.426) | (-1.079) | | | |
| no UCT/Secret | 0.0106 | -0.0385 | | | |
| | (0.266) | (-1.003) | | | |
| UCT/Secret | -0.128*** | -0.0168 | | | |
| | (-3.047) | (-0.438) | | | |
| Constant | 0.648^{***} | 0.249^{***} | | | |
| | (21.54) | (8.211) | | | |
| Observations | 2,284 | 2,284 | | | |
| R-squared | 0.045 | 0.003 | | | |
| p-value of marginal effect [‡] | [0.046] | [0.230] | | | |

Table 11: Treatment effects on demand for agency

Notes: This Table combines observations on all the deferral decisions taken by the wife (column 1) and the husband (column 2) in decision domains A, D, and J. The dependent variable equal 1 if the subject defers, and 0 otherwise. As in Tables 5 to 9, regressors UCT treatment, Secret condition, and UCT x Secret, each corresponds to a different treatment cell. The other regressors are added on top of that. Deferral choices made in domain A are the omitted category. Dummies for domains D and J (Block5) are included. We also include dummies for deferral decisions made for domains A and D in Block2, when deferral either decreases or increases the allocatable budget; the dummy is 1 if deferral is costly and -1 if non-deferral is costly. We also include a dummy equal to -1 if non-deferral is costly in the food and drink game (Block5). The intercept gives the value of the dependent variable in the no-UCT/no-secret treatment cell, domain A, and no cost condition. Each regressor corresponds to a different treatment category with the intercept corresponding to the no-UCT/no-secret category.

 \ddagger is the p-value of the F-test that UCT / Secret = UCT + Secret and tests the marginal contribution of the combined treatments compared to the sum of the contributions.

Robust t-statistics in parentheses, clustered by household. *** p<0.01, ** p<0.05, * p<0.1

 Table 12: Treatment effect on paired deferral rates

| Dependent variable is the proportion of choice pairs in each category | | | | | | | |
|---|---------------|---------------|---------------|---------------|--|--|--|
| Deferral? | W yes, H no | W no, H no | W yes, H yes | W no, H yes | | | |
| UCT/no-secret | 0.005 | 0.035 | -0.035 | -0.006 | | | |
| | (0.141) | (1.142) | (-1.333) | (-0.276) | | | |
| no-UCT/Secret | 0.008 | 0.040 | -0.018 | -0.030 | | | |
| | (0.174) | (0.847) | (-0.533) | (-1.062) | | | |
| UCT/Secret | -0.088* | 0.094^{**} | -0.050** | 0.044 | | | |
| | (-1.754) | (2.372) | (-2.061) | (1.278) | | | |
| Constant | 0.525^{***} | 0.228^{***} | 0.156^{***} | 0.092^{***} | | | |
| | (15.300) | (9.639) | (8.672) | (3.517) | | | |
| Observations | 503 | 503 | 503 | 503 | | | |
| R-squared | 0.012 | 0.011 | 0.006 | 0.017 | | | |
| p-value of marg. effect‡ | [0.061] | [0.720] | [0.937] | [0.021] | | | |

Notes: Each column is the proportion of couples in the described category. \ddagger is the p-value of the F-test that UCT or Secret = UCT + Secret and tests the marginal contribution of the combined treatments compared to the sum of the contributions. Robust t-statistics in parentheses, clustered by session. *** p<0.01, ** p<0.05, * p<0.1

| Regressors: | Wife | Husband |
|-----------------------------|-------------|-----------------|
| UCT/no-Secret | 49.70 | 11.30 |
| | (1.639) | (0.382) |
| no UCT/Secret | 4.964 | 23.36 |
| | (0.169) | (0.618) |
| UCT/Secret | 59.84^{*} | 0.907 |
| | (2.004) | (0.025) |
| Intercept | 1,295*** | $1,\!256^{***}$ |
| | (79.89) | (54.13) |
| Observations | 449 | 449 |
| R-squared | 0.015 | 0.001 |
| p-value of marginal effect‡ | 0.906 | 0.410 |

Table 13: Treatment effects on expected material payoff for self

Notes: The dependent variable is a constructed variable combining the 8 most common payoff vectors at the end of the experiment, namely, **split**A-D and deferA-D. Each of these is randomly drawn with equal probability at the end of the experiment. If **split**A or **split**D is drawn, the material payoffs of the wife and husband are **split** and 2500-**split**, respectively. If **split**B is drawn, the material payoff of the husband is 2500-**split**, and the material payoff of the husband is 0. When one of the DeferA-D decisions is drawn, the outcome vector is **split**A-D if the subject does not defer and **split**A-D of the spouse if the subject defers. The dependent variable is the sum of these 8 equal probability outcomes, divided by 6. Each regressor corresponds to a different treatment cell. The intercept gives the value of the dependent variable in the no-UCT/no-secret treatment cell. Each regressor corresponds to a different treatment cell.

The exact formulas used are:

$$\pi_w = \frac{1}{8} (\operatorname{split} A_w (2 - \operatorname{defer} A_w) + (2500 - \operatorname{split} C_w) (2 - \operatorname{defer} C_w) + \operatorname{split} D_w (2 - \operatorname{defer} D_w) + \operatorname{split} A_h \operatorname{defer} A_w + (2500 - \operatorname{split} C_h) \operatorname{defer} C_w + \operatorname{split} D_h \operatorname{defer} D_w)$$

$$\pi_h = \frac{1}{8} ((2500 - \mathbf{split}A_h)(2 - \mathbf{defer}A_h) + (2500 - \mathbf{split}B_h)(2 - \mathbf{defer}B_h) + (2500 - \mathbf{split}D_h)(2 - \mathbf{defer}D_h) + (2500 - \mathbf{split}A_w)\mathbf{defer}A_h + (2500 - \mathbf{split}B_w)\mathbf{defer}B_h + (2500 - \mathbf{split}D_w)\mathbf{defer}D_h)$$

‡ is the p-value of the F-test that UCT / Secret = UCT + Secret and tests the marginal contribution of the combined treatments compared to the sum of the contributions. Robust t-statistics in parentheses, clustered by session. *** p<0.01, ** p<0.05, * p<0.1

| efficiency type | strong | weak | | | |
|---|---------------|---------------|--|--|--|
| Panel I: Dependent variable is the decision by the wife | | | | | |
| UCT/no-Secret | -0.194*** | -0.173*** | | | |
| | (-2.858) | (-2.883) | | | |
| no UCT/Secret | -0.140* | -0.121* | | | |
| | (-1.779) | (-1.787) | | | |
| UCT/Secret | -0.063 | -0.034 | | | |
| | (-0.849) | (-0.516) | | | |
| Constant | | 0.412^{***} | | | |
| | (7.127) | (8.439) | | | |
| Observations | 392 | 503 | | | |
| p-value of marginal effect‡ | [0.011] | [0.004] | | | |
| Panel II: Dependent variable is the decision by the husband | | | | | |
| UCT/no-Secret | -0.036 | -0.019 | | | |
| | (-0.545) | (-0.327) | | | |
| no UCT/Secret | 0.085 | 0.074 | | | |
| | (1.560) | (1.265) | | | |
| UCT/Secret | 0.086 | 0.088 | | | |
| | (1.513) | (1.512) | | | |
| Constant | 0.238^{***} | 0.267^{***} | | | |
| | (5.251) | (5.605) | | | |
| Observations | 392 | 503 | | | |
| p-value of marginal effect‡ | [0.648] | [0.650] | | | |

Table 14: Treatment effects on efficiency in the production game

Notes: Dependent variable is 1 if the subject chooses the efficient outcome in both games (strong) or the one game presented (weak). Each regressor corresponds to a different treatment cell. The intercept is the value of the dependent variable in the no-UCT/no-secret treatment cell. \ddagger is the p-value of the F-test that UCT / Secret = UCT + Secret and tests the marginal contribution of the combined treatments compared to the sum of the contributions. Robust t-statistics in parentheses, clustered by session. *** p<0.01, ** p<0.05, * p<0.1